

# Cutoff for 1-dim variable speed random walks

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- 1 Notations and reviews
- 2 Randomized products of random walks
- 3 Variable speed random walks
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# Outline

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# Total variation and its mixing time

Consider an irreducible Markov chain on a finite set  $\mathcal{S}$  with transition matrix  $K$  and stationary distribution  $\pi$ .

- The total variation of  $(\mathcal{S}, K, \pi)$  is defined by

$$d_{\text{TV}}(t) := \max_{x \in \mathcal{S}} \|K^t(x, \cdot) - \pi\|_{\text{TV}} = \max_{x \in \mathcal{S}, A \subset \mathcal{S}} \{K^t(x, A) - \pi(A)\}.$$

- The mixing time of the total variation is defined by

$$T_{\text{TV}}(\epsilon) = \min\{t \geq 0 \mid d_{\text{TV}}(t) \leq \epsilon\}, \quad \forall \epsilon \in (0, 1).$$

## Simple random walks with reflecting boundaries

Let  $\mathcal{S} = \{0, 1, \dots, n\}$  and  $K, \pi$  be given by

$$\begin{cases} K(i, i \pm 1) = 1/2, \quad \forall 0 < i < n, \\ K(0, 1) = K(n, n-1) = 1, \end{cases}, \quad \begin{cases} \pi(i) = 1/n, \quad \forall 0 < i < n, \\ \pi(0) = \pi(n) = 1/(2n). \end{cases}$$

Let  $\mathcal{S}' = \mathbb{Z}_{2n}$  and  $K', \pi'$  be given by

$$K'(i, i \pm 1) = 1/2, \quad \pi'(i) = 1/(2n).$$

Through the mapping  $\{i, 2n - i\} \subset \mathcal{S}' \mapsto i \in \mathcal{S}$ , one has

$$d_{\text{TV}}(t) := \max_{x \in \mathcal{S}} \|K^t(x, \cdot) - \pi\|_{\text{TV}} = d'_{\text{TV}}(t) := \max_{y \in \mathcal{S}'} \|(K')^t(y, \cdot) - \pi'\|_{\text{TV}}.$$

# Bounding mixing times using spectral information

Observe that  $K'$  has the following eigenvalues and normalized eigenvectors

$$\beta_i = \cos \frac{\pi i}{n}, \quad \phi_i(j) = \sqrt{2} \cos \frac{\pi ij}{n}, \quad i, j \in \mathcal{S}'.$$

To see an upper bound, set  $d'_2(t) = \max_x \|(K')^t(x, \cdot)/\pi' - 1\|_{\ell^2(\pi')}$ . Then,

$$d'_{\text{TV}}(t) \leq \frac{1}{2} d'_2(t) = \frac{1}{2} \left( \sum_{i=1}^n \beta_i^{2t} \right)^{1/2} = \frac{1}{2} \left( \sum_{i=1}^n \left( \cos \frac{\pi i}{n} \right)^{2t} \right)^{1/2}.$$

For a lower bound, taking  $\phi_1$  as a testing function yields

$$d'_{\text{TV}}(t) = \frac{1}{2} \|(K')^t - \pi'\|_{\ell^\infty(\pi') \rightarrow \ell^\infty(\pi')} \geq \frac{(K')^t(0, \phi_1/\sqrt{2})}{2} = \frac{1}{2} \left( \cos \frac{\pi}{n} \right)^t.$$

# Constant speed random walks

As the random walk  $(S', K', \pi')$  does not converge (in distribution), let's consider one continuous time variant, called the constant speed random walk,  $(S', Q, \pi')$  with infinitesimal generator

$$Q(i, j) = K'(i, j) \quad \forall i \neq j, \quad Q(i, i) = -1 \quad \forall i.$$

By setting  $H_t = e^{tQ}$  and  $d''_{TV}(t) := \max_y \|H_t(y, \cdot) - \pi'\|_{TV}$ , one has

$$\frac{1}{2}e^{-\lambda_1 t} \leq d''_{TV}(t) \leq \frac{1}{2} \left( \sum_{i=1}^n e^{-2\lambda_i t} \right)^{1/2}.$$

where  $\lambda_i = 1 - \cos(\pi i/n)$ . As a result of  $\theta^2/5 \leq 1 - \cos \theta \leq \theta^2/2$  with  $|\theta| \leq \pi$ , there are  $C(\epsilon) > 1$  for  $\epsilon \in (0, 1)$  such that

$$C(\epsilon)^{-1}n^2 \leq T''_{TV}(\epsilon) \leq C(\epsilon)n^2 \quad \text{for } n \text{ large enough.}$$

# Some remarks on classical techniques

## Remarks.

- A similar upper bound can be obtained by the tail probability of a random time, say the coupling time or the strong stationary time.
- A similar lower bound can be derived by the Kolmogorov inequality.
- For multi-dimensional walks, let  $d \in \mathbb{N}$ ,  $\mathcal{S} = (\mathbb{Z}_{2n})^d$ ,  $\pi(x) = (2n)^{-d}$  and

$$Q(x, y) = \frac{1}{2d} \quad \forall |x - y| = \sum_{i=1}^d |x_i - y_i| = 1, \quad Q(x, x) = -1 \quad \forall x.$$

For  $\epsilon \in (0, 1)$ , there are  $C(\epsilon) > 1$  such that

$$C(\epsilon)^{-1}(d \log d)n^2 \leq T_{\text{TV}}(\epsilon) \leq C(\epsilon)(d \log d)n^2 \quad \text{for } n \text{ large enough.}$$



# Cutoffs for Markov chains

A family of continuous time Markov chains,  $\mathcal{F} = (\mathcal{S}_n, Q_n, \pi_n)_{n=1}^{\infty}$ , with mixing times,  $(T_{n,\text{TV}})_{n=1}^{\infty}$ , presents a cutoff in the total variation if

$$\lim_{n \rightarrow \infty} \frac{T_{n,\text{TV}}(\epsilon)}{T_{n,\text{TV}}(1 - \epsilon)} = 1 \quad \forall \epsilon \in (0, 1),$$

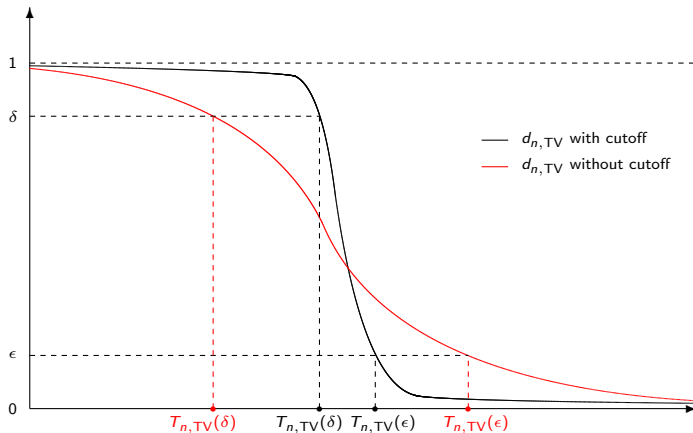
or, equivalently, there is a sequence of positive reals  $(t_n)_{n=1}^{\infty}$  s.t.

$$\lim_{n \rightarrow \infty} d_{n,\text{TV}}(at_n) = \begin{cases} 0 & \text{for } a > 1, \\ 1 & \text{for } 0 < a < 1. \end{cases}$$

The sequence  $(t_n)_{n=1}^{\infty}$  is called a total variation cutoff time for  $\mathcal{F}$ .

**Remark.** For the  $d$ -dimensional constant speed random walks, the family indexed by the sizes of state spaces has no cutoff in the total variation.

# Mixing times and cutoffs



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# Randomized products

For  $n \geq 1$ , let  $\mathcal{X}_n = \{0, 1, \dots, n\}$  and  $R_n$  be a  $Q$ -matrix on  $\mathcal{X}_n$  given by

$$\begin{cases} R_n(i, i+1) = R_n(i, i-1) = 1/2, \quad \forall 0 < i < n, \\ R_n(0, 1) = R_n(n, n-1) = 1, \quad R_n(i, i) = -1. \end{cases}$$

Consider the family  $(\mathcal{S}_n, Q_n, \pi_n)_{n=1}^{\infty}$ , where  $\mathcal{S}_n = \mathcal{X}_n^n$  and

$$Q_n(x, y) = \sum_{i=1}^n \xi_i \delta_{\check{x}_i}(\check{y}_i) R_n(x_i, y_i) \quad \forall x, y \in \mathcal{S}_n,$$

where  $\check{x}_i = (x_1, \dots, x_{i-1}, x_{i+1}, \dots, x_n)$  and  $(\xi_n)_{n=1}^{\infty}$  is a sequence of positive random variables.

# Convergence of randomized products

If  $\nu_n$  is the stationary distribution of  $R_n$ , then

$$\pi_n(x) = \nu_n(x_1) \times \cdots \times \nu_n(x_n).$$

Let  $P_{n,t} = e^{tR_n}$  and  $H_{n,t} = e^{tQ_n}$ . Then, one has

$$H_{n,t}(x, y) = P_{n,t\xi_1}(x_1, y_1) \times \cdots \times P_{n,t\xi_n}(x_n, y_n).$$

## How to analyze the total variation of $(\mathcal{S}_n, Q_n, \pi_n)$ ?

- Coupling for multi-dimensional random walks.
- Spectral analysis with  $(n+1)^n$  eigenvalues and eigenvectors.
- The total variation of  $(\mathcal{X}_n, R_n, \nu_n)$ .

# The Hellinger distance

For any irreducible Markov chain  $(\mathcal{S}, K, \pi)$ , the Hellinger distance is defined by

$$d_H(t) = \sup_{x \in \mathcal{S}} \left( \frac{1}{2} \sum_{y \in \mathcal{S}} \left( \sqrt{K^t(x, y)} - \sqrt{\pi(y)} \right)^2 \right)^{1/2} \in [0, 1].$$

(C. and Kumagai 2016) If  $d_{TV}$  is the total variation of  $(\mathcal{S}, K, \pi)$ , then

$$1 - \sqrt{1 - d_{TV}^2(t)} \leq d_H^2(t) \leq d_{TV}(t).$$

This implies

Cutoff in the total variation  $\Leftrightarrow$  Cutoff in the Hellinger distance.

# Hellinger distances of product chains

Let  $d_{n,H}, D_{n,H}$  be Hellinger distances of  $(\mathcal{X}_n, R_n, \nu_n), (S_n, Q_n, \pi_n)$ . Then,

$$1 - D_{n,H}^2(t) = \prod_{i=1}^n (1 - d_{n,H}^2(\xi_i t)) \leq 1 - d_{n,H}^2(\zeta_n t),$$

where  $\zeta_n = \min\{\xi_1, \dots, \xi_n\}$ , and there is a universal constant  $C > 1$  s.t.

$$C^{-1} e^{-Ct/n^2} \leq d_{n,H}(t) \leq \frac{C e^{-C^{-1}t/n^2}}{1 - e^{-C^{-1}t/n^2}} \quad \forall t > 0, n \geq 1.$$

This implies that, for  $a > 0$ ,

$$-\log(1 - D_{n,H}^2(t)) \asymp \sum_{i=1}^n d_{n,H}^2(\xi_i t) \quad \forall t \geq a \zeta_n^{-1} n^2.$$

## Cutoffs for randomized products

(C. and Kumagai 2017) The family  $(\mathcal{S}_n, Q_n, \pi_n)_{n=1}^{\infty}$  has a cutoff if and only if there is a sequence of positive reals  $(t_n)_{n=1}^{\infty}$  such that

$$\lim_{n \rightarrow \infty} f_n(at_n) = \begin{cases} \infty & \forall a \in (0, 1), \\ 0 & \forall a \in (1, \infty), \end{cases}, \quad f_n(t) = \sum_{i=1}^n e^{-\xi_i t/n^2}. \quad (1)$$

(C. and Saloff-Coste 2010) Let  $(\xi_i^{(n)})_{i=1}^n$  be an order statistics of  $(\xi_i)_{i=1}^n$ . Then, (1) holds if and only if

$$\lim_{n \rightarrow \infty} \xi_1^{(n)} \max_{1 \leq i \leq n} \frac{\log(i+1)}{\xi_i^{(n)}} = \infty.$$



# Some examples

1. For some  $c > 0$ ,  $\mathbb{P}(\xi_n = c, \forall n) = 1$ . In this case, there is a total variation cutoff for  $(\mathcal{S}_n, Q_n, \pi_n)_{n=1}^{\infty}$  and  $T_{n,TV}(\epsilon) \asymp n^2 \log n$ .
2. Suppose that there is an increasing sequence  $(a_n)_{n=1}^{\infty}$  such that  $\xi_n/a_n$  converges a.s. to a positive random variable. Then,  $(\mathcal{S}_n, Q_n, \pi_n)_{n=1}^{\infty}$  has a cutoff with probability one if and only if  $\sup_n \log(n+1)/a_n = \infty$ .
3. If  $(\theta_n)_{n=1}^{\infty}$  is a i.i.d. sequence of positive random variables with finite mean and  $\xi_n = \theta_1 + \dots + \theta_n$ , then, in the setting of **2.**, one has  $a_n = n$  and thus  $(\mathcal{S}_n, Q_n, \pi_n)_{n=1}^{\infty}$  has no cutoff with probability one.

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# Random walks with random hopping times

For  $n \in \mathbb{N}$ , let  $\mathcal{S}_n = \{0, 1, \dots, n\}$  and  $K_n$  be given by

$$K_n(i, i \pm 1) = 1/2 \quad \forall 0 < i < n, \quad K_n(0, 1) = K_n(n, n-1) = 1.$$

Let  $(\xi_n)_{n=1}^{\infty}$  be a sequence of positive random variables and set

$$Q_n(i, j) = \xi_i K_n(i, j) \quad \forall i \neq j, \quad Q_n(i, i) = -\xi_i.$$

The continuous time Markov chain with  $Q$ -matrix  $Q_n$  has stationary distribution

$$\pi_n(i) = \pi_n(0) \frac{2\xi_0}{\xi_i} \quad \forall 0 < i < n, \quad \pi_n(n) = \pi_n(0) \frac{\xi_0}{\xi_n}.$$

# Review of cutoff criterion

(C. and Saloff-Coste, 2015) Consider the family  $\mathcal{F} = (\mathcal{S}_n, Q_n)_{n=1}^{\infty}$ , where  $\mathcal{S}_n = \{0, \dots, n\}$  and  $Q_n$  is a  $Q$ -matrix satisfying  $Q_n(i, i \pm k) = 0$  for  $k > 1$ .

## Notations:

1. Let  $\tau_i^{(n)}$  be the first hitting time to  $i$  of the chain  $(\mathcal{S}_n, Q_n)$ .
2. Let  $M_n$  be a state such that  $\pi_n([0, M_n]) \asymp \pi_n([M_n, n])$ .
3. Set  $t_n = \mathbb{E}_0 \tau_{M_n}^{(n)} \vee \mathbb{E}_n \tau_{M_n}^{(n)}$  and  $b_n^2 = \text{Var}_0 \tau_{M_n}^{(n)} \vee \text{Var}_n \tau_{M_n}^{(n)}$ .

Then, one has

$$\mathcal{F} \text{ has a cutoff in the total variation} \quad \Leftrightarrow \quad b_n = o(t_n).$$

Further,  $t_n$  is the cutoff time.

# Computations of means and variances of hitting times

By writing  $Q_n = (q_{ij}^{(n)})$ , one has

$$\mathbb{E}_i \tau_{i+1}^{(n)} = \frac{\pi_n([0, i])}{\pi_n(i) q_{i,i+1}^{(n)}}, \quad \text{Var}_i \tau_{i+1}^{(n)} = \frac{1}{\pi_n(i) q_{i,i+1}^{(n)}} \sum_{j=0}^i \pi_n(j) (\mathbb{E}_j \tau_i^{(n)} + \mathbb{E}_j \tau_{i+1}^{(n)}).$$

In the case of VSRW, we have

$$\mathbb{E}_0 \tau_i^{(n)} = i \xi_0^{-1} + 2 \sum_{k=1}^{i-1} (\xi_1^{-1} + \dots + \xi_k^{-1}) \asymp \sum_{k=0}^{i-1} (\xi_0^{-1} + \dots + \xi_k^{-1})$$

and

$$\text{Var}_0 \tau_i^{(n)} \asymp \sum_{j=0}^{i-1} \sum_{k=0}^j (\xi_0^{-1} + \dots + \xi_k^{-1})^2.$$

## Cutoffs for variable speed birth and death chains

(C. and Kumagai, 2017, ongoing)

**Assumption:**  $(\xi_n^{-1})_{n=0}^{\infty}$  is an i.i.d. sequence with finite mean, say  $\mu$ .

By the SLLN, we may select  $M_n = \lfloor n/2 \rfloor$  almost surely and obtain

$$\mathbb{E}_0 \tau_{M_n}^{(n)} \asymp \sum_{k=0}^{M_n-1} (\xi_0^{-1} + \dots + \xi_k^{-1}) \sim \frac{\mu n^2}{4} \quad \text{almost surely,}$$

and

$$\text{Var}_0 \tau_{M_n}^{(n)} \asymp \sum_{j=0}^{M_n-1} \sum_{k=0}^j (\xi_0^{-1} + \dots + \xi_k^{-1})^2 \sim \frac{\mu^2 n^4}{192} \quad \text{almost surely.}$$

## Cutoffs for variable speed birth and death chains

By writing

$$\xi_{n-k}^{-1} + \cdots + \xi_n^{-1} = (\xi_0^{-1} + \cdots + \xi_n^{-1}) - (\xi_0^{-1} + \cdots + \xi_{n-k-1}^{-1}),$$

one may apply the same computation as before to achieve

$$\mathbb{E}_n \tau_{M_n-1}^{(n)} \asymp \sum_{k=0}^{n-M_n-1} (\xi_{n-k}^{-1} + \cdots + \xi_n^{-1}) \sim \frac{\mu n^2}{4} \quad \text{almost surely,}$$

and

$$\text{Var}_n \tau_{M_n}^{(n)} \asymp \sum_{j=0}^{n-M_n-1} \sum_{k=0}^j (\xi_{n-k}^{-1} + \cdots + \xi_n^{-1})^2 \sim \frac{\mu^2 n^4}{192} \quad \text{almost surely.}$$

This implies that  $\mathcal{F}$  has no cutoff with probability one.

## Some remarks

- (C. and Saloff-Coste 2013) Let  $\lambda_n$  be the spectral gap of  $Q_n$ , i.e. the smallest non-zero eigenvalue of  $-Q_n$ . Then,

$$\mathcal{F} \text{ has a cutoff} \iff t_n \lambda_n \rightarrow \infty, \quad t_n = \mathbb{E}_0 \tau_{M_n}^{(n)} \vee \mathbb{E}_n \tau_{M_n}^{(n)}.$$

As no subfamily of  $(\mathcal{S}_n, Q_n)_{n=1}^\infty$  presents a cutoff,  $\lambda_n \asymp t_n^{-1} \asymp n^{-2}$  almost surely.

- (M.-F. Chen) Let  $C = \pi_n([0, M_n]) \wedge \pi_n([M_n, n])$  and

$$\ell_n = \max \left\{ \max_{j < M_n} \sum_{k=j}^{M_n-1} \frac{\pi_n([0, j])}{\pi_n(k) q_{k, k+1}^{(n)}}, \max_{j > M_n} \sum_{k=M_n+1}^j \frac{\pi_n([j, n])}{\pi_n(k) q_{k, k-1}^{(n)}} \right\}.$$

Then,

$$\frac{1}{4\ell_n} \leq \lambda_n \leq \frac{2}{C\ell_n}.$$



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# Perspectives of future work

What could we say about:

- the cutoff for the randomized product of random walks with i.i.d.  $(\xi_n)_{n=1}^\infty$ ;
- the cutoff for the one-dimensional VSRW with i.i.d.  $(\xi_n^{-1})_{n=1}^\infty$  and  $\mathbb{E}\xi_1^{-1} = \infty$ ;
- the cutoff for the randomized product of random walks with random hopping times;
- the scaling limit process and its mixing time.

# Reference

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3. Chen, G.-Y. and Kumagai, T. *Cutoffs for product chains*, 2017.
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*Thank you for your attention!*