

Singular perturbation analysis for countable Markov chains

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Outline

- 1** Background
- 2 Singular perturbation for DTMCs
- 3 Singular perturbation for CTMCs

Perturbed DTMC

Let \tilde{X}_n be a time-homogeneous discrete-time Markov chain (DTMC) on a countable state space \mathbb{E} with an **irreducible** stochastic transition matrix \tilde{P} . Suppose that \tilde{P} is **nearly decomposable**, i.e.

$$\tilde{P} = P + \Delta,$$

- ▶ the (small) perturbation matrix Δ satisfies that $\Delta \mathbf{e} = 0$.
- ▶ the unperturbed transition matrix P is **decomposable (reducible)**, the state space \mathbb{E} is decomposed into denumerable irreducible and ergodic classes \mathbb{E}_n for $n \in \hat{\mathbb{E}}$, where the set $\hat{\mathbb{E}} := \{0, 1, \dots, \ell\}$, $0 \leq \ell \leq \infty$. Thus $\mathbb{E} = \bigcup_{n \in \hat{\mathbb{E}}} \mathbb{E}_n$.

The unperturbed transition matrix P can be written as

$$P = \begin{bmatrix} P_0 & 0 & 0 & \cdots \\ 0 & P_1 & 0 & \cdots \\ 0 & 0 & P_2 & \cdots \\ \vdots & \vdots & \vdots & \ddots \end{bmatrix}.$$

When $\ell = 0$, the perturbation is called regular perturbation;
when $\ell \geq 1$, the perturbation is called singular perturbation.

Perturbed CTMC

Let \tilde{X}_t be a time-homogeneous CTMC on a countable space \mathbb{E} with an **irreducible, conservative and possibly unbounded generator \tilde{Q}** . Suppose that \tilde{Q} is **nearly decomposable**, i.e.

$$\tilde{Q} = Q + \Delta,$$

where Δ is the small and Q is decomposable:

$$Q = \begin{pmatrix} Q_0 & 0 & 0 & \cdots \\ 0 & Q_1 & 0 & \cdots \\ 0 & 0 & Q_2 & \cdots \\ \vdots & \vdots & \vdots & \ddots \end{pmatrix}.$$

Similarly, regular perturbation when $\ell = 0$, and singular perturbation when $\ell \geq 1$

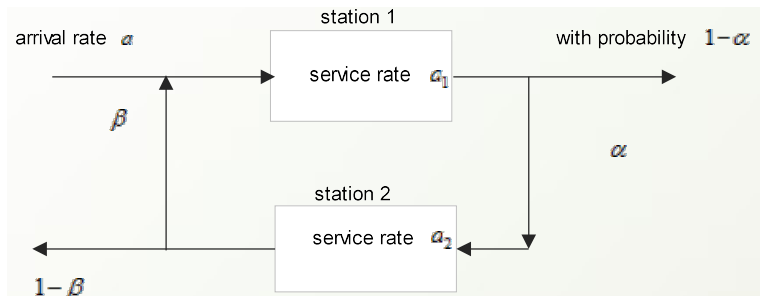
For singular perturbation, the elements of P (Q) and Δ reflect **the strong and weak interactions**, respectively.

A wide range of large-scale systems are distinguished by this feature and hence **are nicely modeled with the help of decomposable Markov chains**, such as

- ▶ Markov decision processes, e.g., Bielecki and Stetlner (1998), Yin and Zhang (2000);
 - ▶ control problems, e.g., Delebecque and Quadrat (1981), Yin and Zhang (2003);
 - ▶ queueing networks, e.g., Latouche and Schweitzer (1995), Yin and Zhang (2008)
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Example 1: a 2d discrete-time queue

Consider a DT queueing system with two stations:



which is modified from [Yin and Zhang \(2008\)](#) by letting that Station 1 has an unlimited room and the arrival customers follow a geometric distribution with a constant parameter.

Let $\tilde{X}_j(n)$ be the number of customers at Station j at time n . Then $(\tilde{X}_2(n), \tilde{X}_1(n))$ is a two-dimensional DTMC on the state space $\mathbb{E} = \bigcup_{n \in \hat{\mathbb{E}}} \hat{\mathbb{E}} \times \mathbb{E}_n$, with $\hat{\mathbb{E}} = \mathbb{E}_n = \{0, 1, 2, \dots\}$.

Its transition matrix \tilde{P} is given by $\tilde{P} = P + \Delta$, where P is decomposable and $P_n, n \geq 1$ and Δ are determined by

$$P_0 = \begin{bmatrix} a & 1-a & 0 & \cdots \\ \mu_1 & \theta_0 & \lambda_1 & \cdots \\ 0 & \mu_1 & \theta_0 & \cdots \\ \vdots & \vdots & \vdots & \ddots \end{bmatrix}, \quad P_n = \begin{bmatrix} c & 1-c & 0 & \cdots \\ \mu_2 & \theta_1 & \lambda_2 & \cdots \\ 0 & \mu_2 & \theta_1 & \cdots \\ \vdots & \vdots & \vdots & \ddots \end{bmatrix}.$$

$$\Delta = \begin{bmatrix} \bar{\Delta}_0 & \bar{\Delta}_1 & 0 & \cdots \\ \Delta_{-1} & \Delta_0 & \Delta_1 & \cdots \\ 0 & \Delta_{-1} & \Delta_0 & \cdots \\ \vdots & \vdots & \vdots & \ddots \end{bmatrix}, \quad \bar{\Delta}_0 = \begin{bmatrix} 0 & 0 & 0 & \cdots \\ 0 & -\hat{\gamma}_{12} & 0 & \cdots \\ 0 & 0 & -\hat{\gamma}_{12} & \cdots \\ \vdots & \vdots & \vdots & \ddots \end{bmatrix},$$

$$\bar{\Delta}_1 = \begin{bmatrix} 0 & 0 & 0 & 0 & \cdots \\ a\hat{\gamma}_{12} & (1-a)\hat{\gamma}_{12} & 0 & 0 & \cdots \\ 0 & a\hat{\gamma}_{12} & (1-a)\hat{\gamma}_{12} & 0 & \cdots \\ \vdots & \vdots & \vdots & \vdots & \ddots \end{bmatrix},$$

$$\Delta_{-1} = \begin{bmatrix} a\gamma_2 & \widehat{\delta}_0 & \overline{\delta}_0 & 0 & 0 & \cdots \\ a\gamma_1\gamma_2 & \delta_1 & \widehat{\delta}_1 & \overline{\delta}_1 & 0 & \cdots \\ 0 & a\gamma_1\gamma_2 & \delta_1 & \widehat{\delta}_1 & \overline{\delta}_1 & \cdots \\ \vdots & \vdots & \vdots & \vdots & \vdots & \ddots \end{bmatrix},$$

$$\Delta_1 = \begin{bmatrix} 0 & 0 & 0 & 0 & \cdots \\ aa_2\widehat{\gamma}_{12} & (1-a)a_2\widehat{\gamma}_{12} & 0 & 0 & \cdots \\ 0 & aa_2\widehat{\gamma}_{12} & (1-a)a_2\widehat{\gamma}_{12} & 0 & \cdots \\ \vdots & \vdots & \vdots & \vdots & \ddots \end{bmatrix},$$

and

$$\Delta_0 = \begin{bmatrix} -\bar{\gamma}_2 & 0 & 0 & 0 & \cdots \\ a\hat{\gamma}_{12}\gamma_2 & \hat{\theta}_1 & (1-a)\hat{\gamma}_{12}\hat{\gamma}_{21} & 0 & \cdots \\ 0 & a\hat{\gamma}_{12}\gamma_2 & \hat{\theta}_1 & (1-a)\hat{\gamma}_{12}\hat{\gamma}_{21} & \cdots \\ \vdots & \vdots & \vdots & \vdots & \ddots \end{bmatrix}.$$

Assume that parameters α and $(1 - a_2)$ are small $\Rightarrow \Delta$ is small.

\Rightarrow the changes of the queue length process and departure process corresponding to Station 2 are relatively slow compared with those corresponding to Station 1, which results in that the transition probability between any two levels is very small.

Example 2: a M M Birth-death process

Consider a Markov modulated state dependent birth-death process.

The environment process is a slowly varying M/M/1 queue, in which customers arrive at this queue according to a Poisson process with rate εq and are served at rate εp . Let $X(t)$ be the number of customers in the queue at time t .

Define a birth-death process $Y(t)$ as follows: if at time t , $X(t) = n$ and $Y(t) = i$, then $Y(t)$ jumps up to $i + 1$ at the birth rate $r_i b_n$ and jumps down to $i - 1$ at the death rate $s_i a_n$ for non-negative functions r_i and s_i on $\{0, 1, 2, \dots\}$.

The process $(X(t), Y(t))$ is a two-dimensional level dependent QBD process with infinitely many levels and phases.

Its generator is $\tilde{Q} = Q + \Delta$, where Q is decomposable with $Q_n, n \geq 0$ given by

$$Q_n = \begin{pmatrix} -r_0 b_n & r_0 b_n & 0 & \cdots \\ s_1 a_n & -r_1 b_n - s_1 a_n & r_1 b_n & \cdots \\ 0 & s_2 a_n & -r_2 b_n - s_2 a_n & \cdots \\ \vdots & \vdots & \vdots & \ddots \end{pmatrix},$$

and Δ has the same form as that of Example 1 with

$$\Delta_1 = \begin{pmatrix} \varepsilon q & 0 & 0 & \cdots \\ 0 & \varepsilon q & 0 & \cdots \\ 0 & 0 & \varepsilon q & \cdots \\ \vdots & \vdots & \vdots & \ddots \end{pmatrix}, \quad \Delta_{-1} = \begin{pmatrix} \varepsilon p & 0 & 0 & \cdots \\ 0 & \varepsilon p & 0 & \cdots \\ 0 & 0 & \varepsilon p & \cdots \\ \vdots & \vdots & \vdots & \ddots \end{pmatrix},$$

where $\bar{\Delta}_1 = \Delta_1$, $\bar{\Delta}_0 = -\bar{\Delta}_1$, $\Delta_0 = -(\Delta_{-1} + \Delta_1)$.

Our goal

Understand how perturbation affect the long-run behavior:

(i) Find conditions on Δ and P_i to ensure that \tilde{P} is positive recurrent with stationary distribution $\tilde{\pi}$, and

$$\tilde{\pi} = \pi^{(0)} \sum_{n=0}^{\infty} U^n,$$

for some matrix U and probability vector $\pi^{(0)}$.

(ii) When (i) is addressed well, we consider the bound on the difference $\tilde{\pi} - \pi^{(0)} = \pi^{(0)} \sum_{n=1}^{\infty} U^n$.

Literature

The singular perturbation problems are investigated by posing the Dobelin condition, see, e.g.

- ◇: Korolyuk and Turbin (1993), Bielecki and Stetlner (1998) for singular perturbation of Markov chains on general measurable state,
- ◇: Yin and Zhang (2002, 2008) for two-time scales singular perturbation.

The Dobelin condition is quite restrictive for DTMCs on an infinitely countable state space, see e.g. Hou and L (2004).

Altman et al. (2004): first adopt geometric ergodicity condition.

- ◇: the bounds are not well investigated,
- ◇: \tilde{P} and P_n are aperiodic,
- ◇: the perturbation Δ is linear (i.e. $\Delta = \varepsilon G$),
- ◇: the generator \tilde{Q} and Q are bounded.

We will investigate these issues by extending the ideas for regularly perturbed MCs, see e.g.

- ◇: Kartashov (1986), Mouhoubi and Aissani (2010), L (2012) for DTMCs,
- ◇: Mitrophanov (2006), Heidergott et al. (2010), L (2015) for CTMCs.

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Notations

Recall that each class P_n is positive recurrent, whose stationary probability vector is π_n and stationary probability matrix is Π_n .

Define the fundamental matrix by $R_n = (I - P_n + \Pi_n)^{-1}$.

Let

$$\Pi = \begin{bmatrix} \Pi_0 & 0 & 0 & \cdots \\ 0 & \Pi_1 & 0 & \cdots \\ 0 & 0 & \Pi_2 & \cdots \\ \vdots & \vdots & \vdots & \ddots \end{bmatrix}, \quad \mathbb{R} = \begin{bmatrix} R_0 & 0 & 0 & \cdots \\ 0 & R_1 & 0 & \cdots \\ 0 & 0 & R_2 & \cdots \\ \vdots & \vdots & \vdots & \ddots \end{bmatrix}.$$

To introduce **the aggregated CTMC** (e.g. Delebecque 1983), define the matrices M and W by

$$M = \begin{bmatrix} \pi_0 & 0 & 0 & \cdots \\ 0 & \pi_1 & 0 & \cdots \\ 0 & 0 & \pi_2 & \cdots \\ \vdots & \vdots & \vdots & \ddots \end{bmatrix}_{\hat{E} \times E}, \quad W = \begin{bmatrix} \mathbf{e}_0 & 0 & 0 & \cdots \\ 0 & \mathbf{e}_1 & 0 & \cdots \\ 0 & 0 & \mathbf{e}_2 & \cdots \\ \vdots & \vdots & \vdots & \ddots \end{bmatrix}_{E \times \hat{E}}.$$

The generator \hat{Q} of the aggregated chain is defined by

$$\hat{Q} = M\Delta W, \quad (\hat{q}_{ij} = \pi_i \Delta_{ij} \mathbf{e}_j, \quad i, j \in \hat{E}),$$

whose stationary distribution and deviation matrix are denoted respectively by $\hat{\pi}$ and $D_{\hat{Q}}$ if they exist.

a key matrix

Introduce a key matrix U as follows

$$U = \Delta(R - \Pi)(I + \Delta W D_{\hat{Q}} M).$$

Remark:

- (i) When $\ell = 0$, $U = \Delta(R - \Pi)$, which plays a key role for regular perturbation analysis.
- (ii) Compared with Altman et al. (2004), this definition permits that \tilde{P} and each class P_n are periodic and the perturbation Δ is general.

\mathbf{v} -norm and Assumption 1

Let $\mathbf{v} \geq 1$. For a measure $\boldsymbol{\mu}$, its \mathbf{v} -norm is defined by $\|\boldsymbol{\mu}\|_{\mathbf{v}} = \sum_{i \in E} |\mu_i| v_i$. For a matrix A , its \mathbf{v} -norm is defined by

$$\|A\|_{\mathbf{v}} = \sup_{i \in E} v_i^{-1} \sum_{j \in E} |A_{ij}| v_j.$$

Assumption 1: Let $\mathbf{v} \geq 1$. Assume that

- (i) all the matrices P , Δ , Π , $\hat{\pi}M$, \mathbb{R} and $WD_{\hat{Q}}M$ have finite \mathbf{v} -norm; and
- (ii) there exists a finite positive integer N and a constant $\delta_N \in (0, 1)$ such that $\delta_N = \|U^N\|_{\mathbf{v}} < 1$.

Main results

Theorem 1: Suppose that Assumption 1 holds. Then

(i) \tilde{P} has a unique stationary distribution, given by

$$\tilde{\pi} = \hat{\pi} M \sum_{k=0}^{\infty} U^k.$$

(ii) For $N = 1$ (the case $N > 1$ can be obtained)

$$\left\| \tilde{\pi} - \hat{\pi} M \sum_{k=0}^m U^k \right\|_{\mathbf{v}} \leq \|\hat{\pi} M\|_{\mathbf{v}} \frac{\|U\|_{\mathbf{v}}^{m+1}}{1 - \|U\|_{\mathbf{v}}}.$$

Remark:

- ▶ Note that for regular perturbation $\hat{\pi}M = \pi$ ($\pi P = \pi$), while for singular perturbation $\hat{\pi}M = (\hat{\pi}_0 \boldsymbol{\pi}_1, \hat{\pi}_1 \boldsymbol{\pi}_1, \dots)$.
- ▶ We call $\sum_{k=0}^m \hat{\pi} M U^k$ the $(m+1)$ -th order approximation of the stationary distribution $\tilde{\pi}$.
- ▶ For regular perturbation, the first approximation $\hat{\pi}M$ (i.e., $m = 0$) becomes the stationary distribution of P , which has caused much concern in the literature.
- ▶ To prove this theorem, we need to make use of **associativity for the multiplications of (possibly negative) infinite matrices**. (i) of Assumption 1 can guarantee the associativity.

We now state **Lyapunov drift conditions** (see Chen (2003)), which are equivalent to geometric and exponential ergodicity, respectively.

D1($\mathbf{V}, \lambda, b, \{i\}$): For a transition matrix P , suppose that there exist a finite vector \mathbf{V} , $\mathbf{V} \geq \mathbf{e}$, some state i and positive numbers $b < \infty, \lambda < 1$ such that

$$P\mathbf{V} \leq \lambda\mathbf{V} + b\mathbb{1}_{\{i\}}.$$

D2($\mathbf{V}, \lambda, b, \{i\}$): For a q -matrix Q , suppose that there exist a finite function \mathbf{V} , $\mathbf{V} \geq \mathbf{e}$, some state i and positive constants $\lambda > 0, b < \infty$ such that

$$Q\mathbf{V} \leq -\lambda\mathbf{V} + b\mathbb{1}_{\{i\}}.$$

Assumption 2: Assume that

- (i) $\mathbf{D1}(\mathbf{V}, \lambda, b, \{i_0(n)\})$ holds uniformly for each P_n ; and
- (ii) $\mathbf{D2}(\hat{\mathbf{V}}, \hat{\lambda}, \hat{b}, \{\hat{i}_0\})$ holds for the generator \hat{Q} .

Theorem 2: Suppose that Assumption 2 holds and $\|\Delta\|_{\mathbf{v}} < \infty$, where $v_{ni} := \hat{V}_n V_i$, $n \in \hat{\mathbb{E}}, i \in \mathbb{E}_n$. Then Assumption 1 holds and

$$\|U\|_{\mathbf{v}} \leq x(1 + y\|\Delta\|_{\mathbf{v}} + z),$$

where x, y, z are given by

$$x = \frac{c_1}{1 - \lambda}, \quad y = \frac{c_2(1 + c_3)^2}{\hat{\lambda}}, \quad z = c_2 c_3$$

with $c_1 := \|\Delta(I - \Pi)\|_{\mathbf{v}}$, $c_2 := \sup_{n \in \hat{\mathbb{E}}} \pi_n(V)$ and $c_3 := \hat{\pi}(\hat{V})$.
 Note that $c_1 \leq (1 + c_2)\|\Delta\|_{\mathbf{v}}$, $c_2 \leq b/(1 - \lambda)$, and $c_3 \leq \hat{b}/\hat{\lambda}$.

Back to Example 1

Assumption-1: Suppose that (i) $a_1 + a\alpha < a + aa_1\alpha$; and (ii) $a_2 + \alpha < 1 + aa_2\alpha$.

▷ Note that (i) of Assumption-1 is equivalent to $\rho := \frac{\lambda_1}{\mu_1} < 1$.

▷ Under (i) of Assumption-1, $\mathbf{D1}(\mathbf{V}, \lambda, b, \{i_0(n)\})$ holds uniformly in n for $i_0(n) = 0$, $\mathbf{V}_j = \left(\sqrt{\frac{1}{\rho}}\right)^j$, $j \geq 0$,

$\lambda = 1 - (\sqrt{\lambda_1} - \sqrt{\mu_1})^2$ and $b = a + (1 - a)\sqrt{\frac{1}{\rho}} - \lambda_0$.

$$\pi_{n,0} = \frac{\frac{a_1}{\rho}}{\frac{a_1}{\rho} + \frac{1}{1-\rho}}, \quad \pi_{n,i} = \frac{\rho^{i-1}}{\frac{a_1}{\rho} + \frac{1}{1-\rho}}, \quad i \geq 1.$$

The generator of the aggregated CTMC \hat{Q} is

$$\hat{Q} = \begin{bmatrix} d_0 & d_1 & 0 & \cdots \\ c_0 & c_1 & c_2 & \cdots \\ 0 & c_0 & c_1 & \cdots \\ \vdots & \vdots & \ddots & \cdots \end{bmatrix},$$

where $d_i = \pi_0 \Delta_{0i} \mathbf{e}$, and $c_i = \pi_0 \Delta_i \mathbf{e}$.

▷ Note that (ii) of Assumption-1 is equivalent to $\hat{\rho} := \frac{c_2}{c_0} < 1$.

▷ When $\hat{\rho} < 1$, **D2**($\hat{\mathbf{V}}, \hat{\lambda}, \hat{b}, \{\hat{i}_0\}$) holds for $\hat{i}_0 = 0$, $\hat{V}_j = \left(\sqrt{\frac{1}{\hat{\rho}}}\right)^j$, $j \geq 0$, $\hat{\lambda} = (\sqrt{c_2} - \sqrt{c_0})^2$, $\hat{b} = d_0 + d_1 \sqrt{\frac{1}{\hat{\rho}}} + \bar{\lambda}_d$.

Define the drift vector \mathbf{v} for the whole space by

$$\mathbf{v}_{ni} = \hat{V}_n V_i = \left(\sqrt{\frac{1}{\hat{\rho}}} \right)^n \left(\sqrt{\frac{1}{\rho}} \right)^i.$$

Assumption-2: Both parameters α and $(1 - a_2)$ are small enough such that $x(1 + y\|\Delta\|_{\mathbf{v}} + z) < 1$.

Under the above two assumptions, by Theorem 2, we have

$$\left\| \tilde{\pi} - \hat{\pi} M \sum_{k=0}^m U^k \right\|_{\mathbf{v}} \leq \frac{[x(1 + y\|\Delta\|_{\mathbf{v}} + z)]^{m+1}}{1 - x(1 + y\|\Delta\|_{\mathbf{v}} + z)} \|\hat{\pi} M\|_{\mathbf{v}},$$

Note that $\|\Delta\|_{\mathbf{v}}$, $\|\hat{\pi} M\|_{\mathbf{v}}$, x , y , and z can be determined by the model parameters.

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Singular perturbation for CTMCs

Let π_n and D_n respectively be the stationary distribution and the deviation matrix corresponding to Q_n whenever they exist.

Let

$$D_Q = \begin{bmatrix} D_0 & 0 & 0 & \cdots \\ 0 & D_1 & 0 & \cdots \\ 0 & 0 & D_2 & \cdots \\ \vdots & \vdots & \vdots & \ddots \end{bmatrix}.$$

The generator \hat{Q} of the aggregated CTMC is defined by

$$\hat{Q} = M\Delta W.$$

Let $\hat{\pi}$ and $D_{\hat{Q}}$ respectively be the stationary distribution and the deviation matrix of \hat{Q} whenever they exist.

To establish the counterpart of the discrete-time perturbation results, define

$$U = \Delta D_Q (I + \Delta W D_{\hat{Q}} M).$$

Assumption 1': Let $\mathbf{v} \geq 1$ be a vector on \mathbb{E} .

- (i) All the matrices Q , Δ , Π , $\hat{\pi}M$, D_Q and $W D_{\hat{Q}} M$ have finite \mathbf{v} -norm;
- (ii) there exists a finite number N and a constant $\delta_N \in (0, 1)$ such that $\delta_N = \|U^N\|_{\mathbf{v}} < 1$.

Theorem 1' Suppose that Assumption 1' holds. Then the same results as that in Theorem 1 hold.

Assumption 2': Assume that

- (i) $\mathbf{D2}(\mathbf{V}, \lambda, b, \{i_0(n)\})$ holds uniformly in n for the same \mathbf{V}, λ and b ; and
- (ii) $\mathbf{D2}(\hat{\mathbf{V}}, \hat{\lambda}, \hat{b}, \{\hat{i}_0\})$ holds for \hat{Q} .

Theorem 2' Suppose that Assumption 2' and $\|\Delta\|_{\mathbf{v}} < \infty$, where $v_{ni} := \hat{V}_n V_i$ for $n \in \hat{\mathbb{E}}, i \in \mathbb{E}_n$. Then (i) of Assumption 1' holds and

$$\|U\|_{\mathbf{v}} \leq x \left[1 + \frac{(1 + \hat{\pi}(\hat{\mathbf{V}}))^2}{\hat{\lambda}} \sup_{n \in \hat{\mathbb{E}}} \pi_n(\mathbf{V}) \|\Delta\|_{\mathbf{v}} \right],$$

where the value of x is given by

$$x = \|\Delta D_Q\|_{\mathbf{v}} \leq \frac{(1 + \sup_{n \in \hat{\mathbb{E}}} \pi_n(\mathbf{V}))^2}{\lambda} \|\Delta\|_{\mathbf{v}}.$$

To prove the above result, one key point is to use the following result

Theorem [L 2015] For a CTMC with a (possibly unbound) Q -matrix. Suppose that $\mathbf{D2}(\mathbf{V}, \lambda, b, \{i\})$ holds. Then we have

$$\|D\|_{\mathbf{v}} \leq \frac{(1 + \pi(\mathbf{V}))^2}{\lambda} \leq \frac{(\lambda + b)^2}{\lambda^3}.$$

Back to Example 2

Assumption: Assume that (i) $\inf_{n \geq 0} (a_n - b_n) > 0$, $\sup_{n \geq 0} b_n < \infty$; and (ii) $p < q$.

(i) of this assumption implies that $\mathbf{D2}(\mathbf{V}, \lambda, b, \{i_0(n)\})$ holds uniformly in n for $i_0(n) = 0$, $V_j = j + 1, j \geq 0$, $\lambda = \frac{\inf_{n \geq 0} (a_n - b_n)}{2}$, and $b = \sup_{n \geq 0} b_n$.

The stationary distribution $\{\pi_n, n \in \hat{\mathbb{E}}\}$, dependent on n , is given by

$$\pi_{n,0} = \frac{1}{1 - \log(1 - b_n/a_n)}, \quad \pi_{n,k} = \frac{1}{k} \left(\frac{b_n}{a_n}\right)^k \pi_{n,0}, \quad k \geq 1.$$

The generator \hat{Q} of the aggregated CTMC is given by

$$\hat{Q} = \begin{pmatrix} -\varepsilon q & \varepsilon q & 0 & \cdots \\ \varepsilon p & -\varepsilon(p+q) & \varepsilon q & \cdots \\ 0 & \varepsilon p & -\varepsilon(p+q) & \cdots \\ \vdots & \vdots & \vdots & \ddots \end{pmatrix}.$$

(ii) of the assumption implies that $\hat{\rho} := \frac{q}{p} < 1$, and

$\mathbf{D2}(\hat{\mathbf{V}}, \hat{\lambda}, \hat{b}, \{\hat{i}_0\})$ holds for $\hat{i}_0 = 0$, $\hat{V}_j = \left(\sqrt{\frac{1}{\hat{\rho}}}\right)^j$, $j \geq 0$,

$\hat{\lambda} = \varepsilon(\sqrt{p} - \sqrt{q})^2$, $\hat{b} = \varepsilon(p - \sqrt{pq})$. Furthermore, we have

$$\hat{\pi}_0 = 1 - \hat{\rho}, \quad \hat{\pi}_n = (1 - \hat{\rho})\hat{\rho}^{n-1}, \quad n \geq 1, \quad \hat{\pi}(\hat{\mathbf{V}}) = 1 + \sqrt{\hat{\rho}}.$$

Define the drift vector \mathbf{v} for the whole space by

$$\mathbf{v}_{ni} = \hat{V}_n V_i = \left(\sqrt{\frac{1}{\hat{\rho}}} \right)^n (i+1), \quad n \geq 0, i \geq 0.$$

Let $c = \sup_{n \geq 0} \pi_n(\mathbf{V})$ and

$$u = \frac{(1+c)^2(\sqrt{p} + \sqrt{q})^2}{\lambda} \left[1 + c(2 + \sqrt{q/p})^2 \frac{(\sqrt{p} + \sqrt{q})^4}{(p-q)^2} \right].$$

Under the above assumption, if $0 < \varepsilon < 1/u$, then we have

$$\left\| \tilde{\pi} - \hat{\pi} M \sum_{k=0}^m U^k \right\|_{\mathbf{v}} \leq c(1 + \sqrt{\hat{\rho}}) \frac{(\varepsilon u)^{m+1}}{1 - \varepsilon u}.$$

Thank you for your attention!