

Limit Theorems for Some Critical Superprocesses

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Outline

References

This talk is based on the following paper:

[1]. Y.-X. Ren, R. Song and R. Zhang: Limit theorems for some critical superprocesses, arXiv:1403.1342.

References

For related works see the following joint papers with R. Song and R. Zhang:

[2]. Central limit theorems for super Ornstein-Uhlenbeck processes. *Acta Appl. Math.* **130** (2014), 9–49.

[3]. Central limit theorems for supercritical branching Markov processes. *J. Funct. Anal.* **266** (2014), 1716-1756.

[4]. Central limit theorems for supercritical branching nonsymmetric Markov processes. [arXiv:1404.0116](https://arxiv.org/abs/1404.0116)

[5]. Central limit theorems for supercritical superprocesses. [arXiv:1310.5410](https://arxiv.org/abs/1310.5410)

[6]. Functional central limit theorems for supercritical superprocesses. Preprint, 2014.

Outline

- 1 Motivation**
- 2 Superprocesses
- 3 Assumptions
- 4 Main Results

For **discrete time critical branching processes** $\{Z(n), n \geq 0\}$, it is known that $P(Z(n) > 0) \rightarrow 0$ as $n \rightarrow \infty$.

Kesten, Ney and Spitzer (1966) proved that if Z has finite second moment, then

$$\lim_{n \rightarrow \infty} nP(Z(n) > 0) = \frac{1}{\sigma^2} \quad (1)$$

and

$$\lim_{n \rightarrow \infty} P\left(\frac{1}{n}Z(n) > \frac{\sigma^2}{2}x \mid Z(n) > 0\right) = e^{-x}, \quad x \geq 0, \quad (2)$$

where σ^2 is the variance of the offspring distribution.

For probabilistic proofs of these results, see Lyons, Pemantle and Peres (1995)

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where σ^2 is the variance of the offspring distribution.

For probabilistic proofs of these results, see Lyons, Pemantle and Peres (1995)

For **continuous time critical branching processes** $\{Z(t), t \geq 0\}$, Athreya and Ney proved in their book (*Branching Processes*, 1972) the following limit theorem: Under the finite second moment condition,

$$\lim_{t \rightarrow \infty} P \left(\frac{1}{t} Z(t) > \frac{\sigma^2}{2} x \mid Z(t) > 0 \right) = e^{-x}, \quad x \geq 0, \quad (3)$$

where σ^2 is a positive constant determined by the branching rate and the variance of the offspring distribution.

For **discrete time multi-type critical branching processes** $\{\mathbf{Z}(n), n \geq 0\}$, Athreya and Ney (1972) gave three limit theorems under the finite second moment condition.
Here $\mathbf{Z}(n) = (Z_1(n), Z_2(n), \dots, Z_d(n))$ (d -type BP).

Let \mathbf{u} and \mathbf{v} be a positive right and left eigenvectors of the mean matrix associated with the eigenvalue 1, respectively.

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Let \mathbf{u} and \mathbf{v} be a positive right and left eigenvectors of the mean matrix associated with the eigenvalue 1, respectively.

(1) $\lim_{n \rightarrow \infty} nP(\mathbf{Z}(n) \neq \mathbf{0} | \mathbf{Z}(0) = \mathbf{i}) = c^{-1}(\mathbf{i} \cdot \mathbf{u})$, where c is a positive constant.

(2) If $\mathbf{w} \cdot \mathbf{v} > 0$, then

$$\lim_{n \rightarrow \infty} P\left(\frac{\mathbf{Z}(n) \cdot \mathbf{w}}{n} > x | \mathbf{Z}(n) > 0\right) = \int_x^\infty f(y) dy, \quad x \geq 0, \quad (4)$$

where $f(y) = \frac{1}{\gamma_1} e^{-y/\gamma_1}$, $y \geq 0$, and γ_1 is a positive constant.

(3) If $\mathbf{w} \cdot \mathbf{v} = 0$, then

$$\lim_{n \rightarrow \infty} P\left(\frac{\mathbf{Z}(n) \cdot \mathbf{w}}{\sqrt{n}} > x | \mathbf{Z}(n) > 0\right) = \int_x^\infty f_2(y) dy, \quad x \in \mathbb{R}, \quad (5)$$

where $f_2(y) = \frac{1}{2\gamma_2} e^{-|y|/\gamma_2}$, $y \in \mathbb{R}$, and γ_2 is a positive constant.

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For **continuous time multi-type critical branching processes**, Athreya and Ney(1974) proved two limit theorems, similar to results (4) and (5) respectively, under the finite second moment condition.

Asmussen and Hering(*Branching Processes*,1983) discussed similar questions for **critical branching Markov processes** $\{Y_t, t \geq 0\}$.

(i) Under some conditions, it was shown that

$$\lim_{t \rightarrow \infty} tP_\nu(\|Y_t\| \neq 0) = c^{-1} \int_E \phi_0(x) \nu(dx).$$

uniformly in ν with ν satisfying $\text{supp}(\nu) = n$ for any integer n , where c is a positive constant and ϕ_0 is the first eigenfunction of the mean semigroup of $\{Y_t, t \geq 0\}$.

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- 2 Superprocesses**
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Superprocesses

E : a locally compact separable metric space.

m : a σ -finite Borel measure on E with full support.

∂ : a separate point not contained in E . ∂ will be interpreted as the cemetery point.

$\xi = \{\xi_t, \Pi_x\}$: a Hunt process on E .

$\zeta := \inf\{t > 0 : \xi_t = \partial\}$ is the lifetime of ξ .

$\{P_t : t \geq 0\}$: the semigroup of ξ .

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The superprocess $X = \{X_t : t \geq 0\}$ we are going to work with is determined by **three parameters**:

- (i) **a spatial motion** $\xi = \{\xi_t, \Pi_x\}$ on E ,
- (ii) **a branching rate function** $\beta(x)$ on E which is a non-negative bounded measurable function.
- (iii) **a branching mechanism** φ of the form

$$\varphi(x, z) = -a(x)z + b(x)z^2 + \int_{(0, +\infty)} (e^{-zy} - 1 + zy)n(x, dy), \quad x \in E, z > 0, \quad (6)$$

where $a \in \mathcal{B}_b(E)$, $b \in \mathcal{B}_b^+(E)$ and n is a kernel from E to $(0, \infty)$ satisfying

$$\sup_{x \in E} \int_{(0, +\infty)} y^2 n(x, dy) < \infty. \quad (7)$$

$\mathcal{M}_F(E)$ denote the space of finite measures on E .
 $\langle f, \mu \rangle := \int_E f(x) \mu(dx)$ and $\|\mu\| := \langle \mathbf{1}, \mu \rangle = \mu(E)$.

The superprocess X is a Markov process taking values in $\mathcal{M}_F(E)$.
 For any $\mu \in \mathcal{M}_F(E)$, we denote the law of X with initial configuration μ by \mathbb{P}_μ . Then for every $f \in \mathcal{B}_b^+(E)$ and $\mu \in \mathcal{M}_F(E)$,

$$-\log \mathbb{P}_\mu \left(e^{-\langle f, X_t \rangle} \right) = \langle u_f(t, \cdot), \mu \rangle, \quad (8)$$

where $u_f(t, x)$ is the unique positive solution to the equation

$$u_f(t, x) + \Pi_x \int_0^{t \wedge \zeta} \varphi(\xi_s, u_f(t-s, \xi_s)) \beta(\xi_s) ds = \Pi_x f(\xi_t), \quad (9)$$

Define

$$\alpha(x) := \beta(x)a(x) \quad \text{and} \quad A(x) := \beta(x) \left(2b(x) + \int_0^\infty y^2 n(x, dy) \right).$$

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For any $f \in \mathcal{B}_b(E)$ and $(t, x) \in (0, \infty) \times E$, define

$$T_t f(x) := \Pi_x \left[e^{\int_0^t \alpha(\xi_s) ds} f(\xi_t) \right].$$

First moment: For any $f \in \mathcal{B}_b(E)$,

$$\mathbb{P}_\mu \langle f, X_t \rangle = \langle T_t f, \mu \rangle.$$

Second moment: For any $f \in \mathcal{B}_b(E)$,

$$\text{Var}_\mu \langle f, X_t \rangle = \langle \text{Var}_\delta \langle f, X_t \rangle, \mu \rangle = \int_E \int_0^t T_s [A(T_{t-s} f)^2](x) ds \mu(dx), \quad (10)$$

where Var_μ stands for the variance under \mathbb{P}_μ .

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- 1 Motivation
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Assumptions on the spatial process

We assume that there exists a family of continuous strictly positive functions $\{p_t(x, y) : t > 0\}$ on $E \times E$ such that

$$P_t f(x) = \int_E p(t, x, y) f(y) m(dy).$$

Define

$$a_t(x) := \int_E p(t, x, y)^2 m(dy), \quad \hat{a}_t(x) := \int_E p(t, y, x)^2 m(dy).$$

Assumption 1

- (i) For any $t > 0$, $\int_E p(t, x, y) m(dx) \leq 1$.
- (ii) For any $t > 0$, we have $a_t(x), \hat{a}_t(x) \in L^1(E, m(dx))$
Moreover, the functions $x \rightarrow a_t(x)$ and $x \rightarrow \hat{a}_t(x)$ are continuous on E .

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Moreover, the functions $x \rightarrow a_t(x)$ and $x \rightarrow \hat{a}_t(x)$ are continuous on E .

One can check that there exists a family $\{q_t(x, y) : t > 0\}$ of continuous strictly positive symmetric functions on $E \times E$ such that

$$T_t f(x) = \Pi_x \left[e^{\int_0^t \alpha(\xi_s) ds} f(\xi_t) \right] = \int_E q_t(x, y) f(y) m(dy).$$

Let $\{\widehat{T}_t, t > 0\}$ be the adjoint operators on $L^2(E, m)$ of $\{T_t, t > 0\}$, that is, for $f, g \in L^2(E, m)$,

$$\int_E f(x) T_t g(x) m(dx) = \int_E g(x) \widehat{T}_t f(x) m(dx)$$

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It follows from (i) above that, for any $t > 0$, T_t is a Hilbert-Schmidt operator and thus a compact operator. Let L and \widehat{L} be the infinitesimal generators of the semigroups $\{T_t\}$ and $\{\widehat{T}_t\}$ in $L^2(E, m)$ respectively.

Define $\lambda_0 := \sup \Re(\sigma(L)) = \sup \Re(\sigma(\widehat{L}))$.

By Jentzsch's theorem, λ_0 is an eigenvalue of multiplicity 1 for both L and \widehat{L} .

Assume that ϕ_0 and ψ_0 are the eigenfunctions of L and \widehat{L} respectively associated with λ_0 . ψ_0 and ϕ_0 can be chosen to be continuous and strictly positive satisfying $\|\phi_0\|_2 = 1$ and $\langle \phi_0, \psi_0 \rangle_m = 1$.

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More Assumptions

Assumption 2

- (i) ϕ_0 is bounded.
- (ii) The semigroup $\{T_t, t > 0\}$ is intrinsically ultracontractive, that is, there exists $c_t > 0$ such that

$$q(t, x, y) \leq c_t \phi_0(x) \psi_0(y). \quad (11)$$

Assumption 3 The superprocess is critical: $\lambda_0 = 0$.

Assumption 4 Define $q_t(x) := \mathbb{P}_{\delta_x}(\|X_t\| = 0)$. We also assume that There exists $t_0 > 0$ such that,

$$\inf_{x \in E} q_{t_0}(x) > 0. \quad (12)$$

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Remarks on our assumptions

In Liu, Ren and Song (2011), quite a few examples of Hunt processes satisfying Assumptions 1 and 3 were given.

If E consists of finitely many points, and $\xi = \{\xi_t : t \geq 0\}$ is a conservative irreducible Markov process on E , then ξ satisfies the Assumptions 1 and 3 for some finite measure m on E with full support. So, as special cases, our results give the analogs of the results of Athreya and Ney (1974) for critical super-Markov chains.

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Main Result

Theorem 1 For any non-zero $\mu \in \mathcal{M}_F(E)$,

$$\lim_{t \rightarrow \infty} t \mathbb{P}_\mu (\|X_t\| \neq 0) = \nu^{-1} \langle \phi_0, \mu \rangle. \quad (13)$$

Main Result

Define $\mathbb{P}_{t,\mu}(\cdot) := \mathbb{P}_\mu(\cdot \mid \|X_t\| \neq 0)$.

Assume that $Y_t, t > 0$, and Y are random variables on (Ω, \mathcal{G}) . We write

$$Y_t|_{\mathbb{P}_{t,\mu}} \rightarrow Y \quad \text{in probability,}$$

if, for any $\epsilon > 0$,

$$\lim_{t \rightarrow \infty} \mathbb{P}_{t,\mu}(|Y_t - Y| \geq \epsilon) = 0.$$

Suppose that Z is a random variable on a probability space $(\tilde{\Omega}, \tilde{\mathcal{G}}, P)$, we write

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if, for all $a \in \mathbb{R}$ with $P(Z = a) = 0$,

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Define

$$\nu := \frac{1}{2} \langle A(\phi_0)^2, \psi_0 \rangle_m. \quad (14)$$

It is easy to see that $0 < \nu < \infty$. Define

$$\mathcal{C}_p := \{f \in \mathcal{B}(E) : \langle |f|^p, \psi_0 \rangle_m < \infty\}.$$

Theorem 2 If $f \in \mathcal{C}_2$ then, for any non-zero $\mu \in \mathcal{M}_F(E)$, we have

$$t^{-1} \langle f, X_t \rangle |_{\mathbb{P}_{t,\mu}} \xrightarrow{d} \langle f, \psi_0 \rangle_m W, \quad (15)$$

where W is an exponential random variable with parameter $1/\nu$. In particular, we have

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Remark Our assumptions imply that $1 \in \mathcal{C}_2$. Thus the limit result above implies that

$$t^{-1} \langle 1, X_t \rangle |_{\mathbb{P}_{t,\mu}} \xrightarrow{d} \langle 1, \psi_0 \rangle_m W,$$

which says that, conditioned on no-extinction at time t , the growth rate of the total mass $\langle 1, X_t \rangle$ is t as $t \rightarrow \infty$.

Note that, when $\langle f, \psi_0 \rangle_m = 0$, $t^{-1} \langle f, X_t \rangle |_{\mathbb{P}_{t,\mu}} \rightarrow 0$ in probability. Therefore it is natural to consider central limit type theorems for $\langle f, X_t \rangle$.

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$$\sigma_f^2 = \int_0^\infty \langle A(T_s f)^2, \psi_0 \rangle_m ds. \quad (17)$$

Theorem 3

Suppose that $f \in C_2$ and $\langle f, \psi_0 \rangle_m = 0$, then we have, $\sigma_f^2 < \infty$ and for any non-zero $\mu \in \mathcal{M}_F(E)$,

$$\left(t^{-1} \langle \phi_0, X_t \rangle, t^{-1/2} \langle f, X_t \rangle \right) |_{\mathbb{P}_{t,\mu}} \xrightarrow{d} \left(W, G(f) \sqrt{W} \right), \quad (18)$$

where $G(f) \sim \mathcal{N}(0, \sigma_f^2)$ is a normal random variable and W is the random variable defined in Theorem 2. Moreover, W and $G(f)$ are independent.

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As a consequence of Theorem 3, we immediately get the following central limit theorem.

Corollary Suppose that $f \in \mathcal{C}_2$ and $\langle f, \psi_0 \rangle_m = 0$, then we have, $\sigma_f^2 < \infty$ and for any non-zero $\mu \in \mathcal{M}_F(E)$,

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Sketch of the proof of Theorem 3

We need to consider the limit of the following \mathbb{R}^2 -valued random variable:

$$U_1(t) := \left(t^{-1} \langle \phi_0, X_t \rangle, t^{-1/2} \langle f, X_t \rangle \right).$$

Which is equivalent to consider the limit of

$$U_1(s+t) = \left((t+s)^{-1} \langle \phi_0, X_{t+s} \rangle, (t+s)^{-1/2} \langle f, X_{t+s} \rangle \right) \quad \text{as } t \rightarrow \infty.$$

First, we consider

$U_2(s, t) = \left(t^{-1} \langle \phi_0, X_t \rangle, t^{-1/2} (\langle f, X_{s+t} \rangle - \langle T_s f, X_t \rangle) \right)$. We prove that

$$U_2(s, t) |_{\mathbb{P}_{t, \mu}} \xrightarrow{d} \left(W, \sqrt{W} G_1(s) \right), \quad \text{as } t \rightarrow \infty, \quad (20)$$

where $G_1(s) \sim \mathcal{N}(0, \sigma_f^2(s))$ with $\sigma_f^2(s) = \langle \text{Var}_\delta \langle f, X_s \rangle, \psi_0 \rangle_m$ and W is the random variable defined in Theorem 3.

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The characteristic function of $U_2(s, t)$ is

$$\begin{aligned}
 & \mathbb{P}_{t,\mu}(\exp\{i\theta_1 t^{-1} \langle \phi_0, X_t \rangle + i\theta_2 t^{-1/2} (\langle f, X_{s+t} \rangle - \langle T_s f, X_t \rangle)\}) \\
 = & \mathbb{P}_{t,\mu} \left(\exp\{i\theta_1 t^{-1} \langle \phi_0, X_t \rangle - i\theta_2 t^{-1/2} \langle T_s f, X_t \rangle + \right. \\
 & \left. \langle \log \mathbb{P}_\delta. \exp\{-i\theta_2 t^{-1/2} \langle f, X_s \rangle\}, X_t \rangle\} \right) \\
 = & \mathbb{P}_{t,\mu} \left(\exp\{i\theta_1 t^{-1} \langle \phi_0, X_t \rangle \right. \\
 & \left. + \int_E \int_{\mathbb{D}} \left(e^{i\theta_2 t^{-1/2} \langle f, \omega_s \rangle} - 1 - i\theta_2 t^{-1/2} \langle f, \omega_s \rangle \right) \mathbb{N}_x(d\omega) X_t(dx) \right),
 \end{aligned}$$

where $\mathbb{P}_{\delta_x} \longleftrightarrow \mathbb{N}_x$ for each $x \in E$. For the definition of \mathbb{N}_x , see Z. Li's book (Measure-valued Branching Markov Processes, 2011)

Thank you!