# Density convergence for some nonlinear Gaussian stationary sequences

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Joint work with

David Nualart (University of Kansas)

Samy Tindel (University Nancy)

Fangjun Xu (East China Normal University)

Based on a joint work with

David Nualart, Samy Tindel, Fangjun Xu

Density convergence in the Breuer-Major theorem for nonlinear Gaussian stationary sequences.

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To appear in Bernoulli.

# Outline

- 1. Motivation
- 2. Main results
- 3. Idea of the proof

# 1. Motivation

Central limit theorem:

Let  $X_1, \dots, X_n$  be independent, identically distributed random variables with mean *m* and variance  $\sigma^2$ .

$$F_n := \sqrt{n} \left( \frac{X_1 + \dots + X_n}{n} - \mathbb{E}(X_1) \right) \longrightarrow \mathcal{N}(0, \sigma^2)$$
$$\frac{X_1 + \dots + X_n}{n} - \mathbb{E}(X_1) \approx \frac{\xi}{\sqrt{n}}, \quad \text{where} \quad \xi \sim \mathcal{N}(0, \sigma^2).$$

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The above convergence is in the sense of distribution

$$F_n \rightarrow N(0, \sigma^2)$$
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$$P(F_n \leq a) 
ightarrow \int_{-\infty}^{a} \phi_{\sigma}(x) dx \quad orall \quad a \in \mathbb{R}$$

where 
$$\phi_{\sigma}(x) = rac{1}{\sqrt{2\pi}\sigma}e^{-rac{x^2}{2\sigma^2}}$$
.

Other examples of multiple Itô integral  $F_n$ 

$$F_n = \int_{[0,T]^q} f_n(t_1,\cdots,t_q) dB_{t_1}\cdots dB_{t_q},$$

where q is an fixed positive integer,  $(B_t, t \ge 0)$  is a standard Brownian motion,  $f_n$  is a sequence of deterministic functions such that

$$\int_{[0,T]^q} f_n^2(t_1,\cdots,t_q) dt_1\cdots dt_q$$

#### Theorem

If  $F_n = I_q(f_n)$ , then the following are equivalent: (i)  $\lim_{n\to\infty} \mathbb{E}[F_n^4] = 3\sigma^4$ , (ii) For all  $1 \le r \le q-1$ ,  $\lim_{n\to\infty} \|f_n \otimes_r f_n\|_{H^{\otimes 2(q-r)}} = 0$ , (iii)  $\|DF_n\|_H^2 \to q\sigma^2$  in  $L^2(\Omega)$  as  $n \to \infty$ . (iv)  $F_n$  converges in distribution to the normal law  $N(0, \sigma^2)$  as  $n \to \infty$ .

Nualart, D.; Peccati, G. Central limit theorems for sequences of multiple stochastic integrals.

Ann. Probab. 33 (2005), 177-93.

Nualart, D.; Ortiz-Latorre, S.

Central limit theorems for multiple stochastic integrals and Malliavin calculus.

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Stochastic Process. Appl. 118 (2008), 614-628.

Let  $X = \{X_k; k = 0, 1, 2, \dots\}$  be a centered Gaussian stationary sequence with unit variance. For all v, we set

$$\rho(\mathbf{v}) = \mathbb{E}[X_0 X_{|\mathbf{v}|}]$$

Let  $\gamma$  be the standard Gaussian probability measure and  $f \in L^2(\gamma)$  be a fixed deterministic function such that  $\mathbb{E}[f(X_1)] = 0$ .

$$f(x) = \sum_{j=d}^{\infty} a_j H_j(x),$$

with  $a_d \neq 0$  and  $d \geq 2$ .

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with  $a_d \neq 0$  and  $d \geq 2$ . Define  $V_n = \frac{1}{\sqrt{n}} \sum_{k=0}^{n-1} f(X_k)$ .

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with  $a_d \neq 0$  and  $d \geq 2$ . Define  $V_n = \frac{1}{\sqrt{n}} \sum_{k=0}^{n-1} f(X_k)$ . The Breuer-Major Theorem Suppose that  $\sum_{\nu=-\infty}^{\infty} |\rho(\nu)|^d < \infty$  and suppose

$$\sigma^2 = \sum_{j=d}^{\infty} j! a_j^2 \sum_{v=-\infty}^{\infty} \rho(v)^j$$

is in  $(0,\infty)$ . Then we have

$$V_n \xrightarrow{\text{Law}} N(0, \sigma^2)$$

P. Breuer and P. Major Central limit theorems for non-linear functionals of Gaussian fields. *J. Mult. Anal.* **13** (1983), 425–441.

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#### Corollary

Consider  $2 \le d \le q < \infty$  and a family of real numbers  $\{a_j; j = d, \ldots, q\}$ . Let  $H_j$  be the jth order Hermite polynomial, and assume that  $\sigma^2 \in (0, \infty)$ , where  $\sigma^2 \equiv \sum_{j=d}^q j! a_j^2 \sum_{v \in \mathbb{Z}} \rho(v)^j$ . Set

$$V_n^{d,q} = \frac{1}{\sqrt{n}} \sum_{k=0}^{n-1} \sum_{j=d}^{q} a_j H_j(X_k).$$

Then  $V_n^{d,q} \xrightarrow{\text{Law}} \mathcal{N}(0,\sigma^2)$  as n tends to infinity. In particular, we have:

$$\lim_{n\to 0} \mathbf{E}\left[\left(V_n^{d,q}\right)^4\right] = 3\sigma^4.$$

There are some development along this direction.

Convergence in density of multiple integrals

Convergence in density of multiple integrals Are there  $f_n(x)$  such that

$$P(F_n \le a) = \int_{-\infty}^a f_n(x) dx$$

and

$$f_n(x) \longrightarrow \phi_\sigma(x)?$$

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## 2. Main result

Let  $X = \{X_k, k \ge 0\}$  be a Gaussian stationary sequence whose spectral density  $f_{\rho}(\lambda) = \frac{1}{2\pi} \sum_{\nu=-\infty}^{\infty} \rho(\nu) e^{i\nu\lambda}$  satisfies  $f_{\rho} \in L^{1/2}([-\pi, \pi])$  and  $\log(f_{\rho}) \in L^{1}([-\pi, \pi])$ , where for all  $\nu$ , we set

 $\rho(\mathbf{v}) = \mathbb{E}[X_0 X_{|\mathbf{v}|}]$ 

## Let

$$V_n^{d,q} = rac{1}{\sqrt{n}} \sum_{k=0}^{n-1} \sum_{j=d}^{q} a_j H_j(X_k),$$

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where  $d \ge 2$ . Assume



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Let

$$V_n^{d,q} = \frac{1}{\sqrt{n}} \sum_{k=0}^{n-1} \sum_{j=d}^{q} a_j H_j(X_k),$$

where  $d \ge 2$ . Assume

$$\sigma^2 \equiv \sum_{j=d}^q j! a_j^2 \sum_{\mathbf{v} \in \mathbb{Z}} \rho(\mathbf{v})^j \in (0,\infty).$$

Then for any  $p \ge 1$ , there exists  $n_0$  such that

$$\sup_{n\geq n_0} \mathbf{E}\left[\|DV_n^{d,q}\|^{-p}\right] < \infty.$$

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In the case of a fixed Wiener chaos we can obtain the following consequence.

#### Corollary

Under the conditions of above theorem, if q = d, and we define  $F_n = V_n^{d,d}/\sigma_n$ , where  $\sigma_n^2 = \mathbf{E}[(V_n^{d,d})^2]$ , then, for all  $m \ge 0$  there exists an  $n_0$  (depending on m) such that

$$\sup_{n\geq n_0}\sup_{x\in\mathbb{R}}|p_{F_n}^{(m)}(x)-\phi^{(m)}(x)|\leq c_m\sqrt{\mathbf{E}[F_n^4]-3}.$$

In the case  $q \neq d$ ,

Corollary

Under the conditions of the above theorem, if we define  $F_n = V_n^{d,q} / \sigma_n$ , where  $\sigma_n^2 = \mathbf{E}[(V_n^{d,q})^2]$ , then, for all  $m \ge 0$  we have

$$\lim_{n\to\infty}\sup_{x\in\mathbb{R}}|p_{F_n}^{(m)}(x)-\phi^{(m)}x)|=0.$$

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## Discussion of the hypothesis

$$f_
ho(\lambda) = rac{1}{2\pi} \sum_{k \in \mathbb{Z}} 
ho(k) \, e^{\imath k \lambda}, \qquad \lambda \in [-\pi,\pi].$$

We assume that

 $f_
ho\in L^{1/2}([-\pi,\pi])$  and  $\log(f_
ho)\in L^1([-\pi,\pi]).$ 

This condition  $\log(f_{\rho}) \in L^1([-\pi, \pi])$  is referred to as *purely nondeterministic* property in the literature.

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## Proposition

Let  $\rho$  be the covariance function of X. We have the following statements.

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(i) If  $\rho \in \ell^1$ , then the spectral density  $f_\rho$  exists and is a nonnegative  $L^2$  function defined on  $[-\pi, \pi]$ . Then the condition is thus fulfilled.

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### Proposition

Let  $\rho$  be the covariance function of X. We have the following statements.

(i) If  $\rho \in \ell^1$ , then the spectral density  $f_{\rho}$  exists and is a nonnegative  $L^2$  function defined on  $[-\pi, \pi]$ . Then the condition is thus fulfilled.

(ii) If  $\lim_{k\to\infty} |k|^{\alpha}\rho(k) = c_{\rho}$  for some  $\alpha \in (0, 1)$  and some positive constant  $c_{\rho}$ , then the spectral density exists, is strictly positive almost everywhere and satisfies  $\lim_{\lambda\to 0} |\lambda|^{1-\alpha} f_{\rho}(\lambda) = c_{f}$ . In particular, the condition is satisfied.

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## Example 1

Gaussian autoregressive fractionally integrated moving-average (Gaussian ARFIMA) processes. Denote by *B* the one lag backward operator ( $BX_k = X_{k-1}$ ). Let  $\phi(z)$  and  $\theta(z)$  be two polynomials which have no common zeros and such that the zeros of  $\phi$  lie outside the closed unit disk  $\{z, |z| \le 1\}$ . Suppose that  $X_k$  is given by

$$\phi(B)X_k = (\mathrm{Id} - B)^{-d}\theta(B)w_k\,,$$

where -1 < d < 1/2, and where the operator  $(Id - B)^{-d}$  is defined by:

$$(\mathrm{Id}-B)^{-d} = \sum_{j=1}^{\infty} \eta_j B^j \quad ext{with} \quad \eta_j = rac{\Gamma(d+j)}{\Gamma(j+1)\Gamma(d)} \, .$$

The sequence  $(w_k)_{k\in\mathbb{Z}}$  is a discrete Gaussian white noise. It is well-known that under the above conditions,  $\{X_k, k \in \mathbb{N}\}$  admits a spectral density whose exact expression is:

$$f(\lambda) = rac{1}{2\pi} \left[ 2\sinrac{\lambda}{2} 
ight]^{-2d} rac{| heta(e^{-i\lambda})|^2}{|\phi(e^{-i\lambda})|^2}.$$

It is thus readily checked that the conditions are satisfied, and hence  $X_k$  has a causal representation.

## Example 2

Our second example is the fractional Gaussian noise. Let  $\{B_t, t \ge 0\}$  be a fractional Brownian motion of Hurst parameter  $H \in (0, 1)$ . Then  $\{X_k = B_{k+1} - B_k, k \in \mathbb{N} \cup \{0\}\}$  is a stationary Gaussian process with correlation

$$\rho(k) = \frac{1}{2} \left[ (k+1)^{2H} - 2k^{2H} - (k-1)^{2H} \right]$$

Its spectral density is:

$$f(\lambda) = \frac{1}{2\pi} \sum_{k=-\infty}^{\infty} \rho(|k|) e^{i\lambda} = 2c_f(1-\cos(\lambda)) \sum_{j=-\infty}^{\infty} |2\pi j + \lambda|^{-2H-1},$$

where  $c_f = (2\pi)^{-1} \sin(\pi H) \Gamma(2H + 1)$ .

If  $H \leq 1/2$ , it is clear that  $\sum_{k=-\infty}^{\infty} |\rho(|k|)| < \infty$ . This implies

$$\sup_{\lambda\in [-\pi,\pi]} |f(\lambda)| <\infty$$
 .

Thus  $f \in L^{1/2}$ . If 1/2 < H < 1, then  $0 \le f(\lambda) \le 2c_f(1 - \cos(\lambda))|\lambda|^{-2H-1} + 2c_f \sum_{j \ne 0} |2\pi j + \lambda|^{-2H-1}$ .

The first term is in  $L^1$  since H < 1. When  $j \neq 0$ ,  $\int_{-\pi}^{\pi} |2\pi j + \lambda|^{-2H-1} d\lambda \leq C j^{-2H}$  for some positive constant C. Thus  $\int_{-\pi}^{\pi} \sum_{j\neq 0} |2\pi j + \lambda|^{-2H-1} d\lambda < \infty$ , owing to the fact that H > 1/2. Therefore, we have  $f \in L^1$  and hence  $f \in L^{1/2}$ . Summarizing we have  $f \in L^{1/2}$  for all  $H \in (0, 1)$ . This also implies  $\log^+ f(\lambda) \in L^1$ . To see  $\log^- f(\lambda) \in L^1$ , we notice that

$$f(\lambda) \ge 2c_f(1 - \cos(\lambda))|\lambda|^{-2H-1}$$

So  $\log^{-} f(\lambda) \leq C + \left| \log \left[ (1 - \cos(\lambda)) |\lambda|^{-2H-1} \right] \right|$  which is in  $L^{1}$ . In conclusion, the sequence X satisfies Hypothesis.

# 3. Idea of the proof

# Causal representation

#### Proposition

Let X be a Gaussian stationary sequence satisfying the hypothesis. Then for each  $k \in \mathbb{N} \cup \{0\}$  the random variable  $X_k$  can be decomposed as

$$X_k = \sum_{j \ge 0} \psi_j \, w_{k-j},$$

where  $(w_k)_{k\in\mathbb{Z}}$  is a discrete Gaussian white noise and the coefficients  $\psi_j$  are deterministic. With a slight abuse of notation, extend the sequence  $\psi$  to  $(\psi_j)_{j\in\mathbb{Z}}$  by setting  $\psi_{-j} = 0$  for  $j \ge 0$ .

Then one can choose  $\psi$  such that it enjoys the following properties: (i) The sequence  $\psi$  admits a spectral density  $f_{\psi}$  such that  $f_{\psi} = \frac{f_{\rho}^{1/2}}{2\pi}$ . (ii) In particular,  $\psi_0 = \frac{1}{2\pi} \int_{-\pi}^{\pi} f_{\rho}^{1/2}(\lambda) d\lambda$  and  $\psi_0 > 0$ . (iii) For all  $k_1, k_2 \in \mathbb{N}$  we have  $\rho(k_1 - k_2) = \sum_{l=-\infty}^{k_1 \wedge k_2} \psi_{k_1 - l} \psi_{k_2 - l}$ .

## Malliavin calculus

Let  $\mathfrak{H}$  be a real separable Hilbert space with inner product  $\langle \cdot, \cdot \rangle_{\mathfrak{H}}$ . The norm of  $\mathfrak{H}$  will be denoted by  $\|\cdot\| = \|\cdot\|_{\mathfrak{H}}$ . Recall that we call *isonormal Gaussian process* over  $\mathfrak{H}$  any centered Gaussian family  $W = \{W(h) : h \in \mathfrak{H}\}$ , defined on a probability space  $(\Omega, \mathcal{F}, \mathbf{P})$  and such that  $\mathbf{E}[W(h)W(g)] = \langle h, g \rangle_{\mathfrak{H}}$  for every  $h, g \in \mathfrak{H}$ . In our application the underlying Gaussian family will be a discrete Gaussian white noise  $(w_k)_{k \in \mathbb{Z}}$ . The space  $\mathfrak{H}$  is given here by  $\mathfrak{H} = \ell^2(\mathbb{Z})$  (the space of square integrable sequences indexed by  $\mathbb{Z}$ )

equipped with its natural inner product. Set  $\{\varepsilon^j; j \in \mathbb{Z}\}$  for the canonical basis of  $\ell^2(\mathbb{Z})$ , that is  $\varepsilon^j_k = \delta_j(k)$ . We thus identify  $w_j$  with  $W(\varepsilon^j)$ . Assume from now on that our underlying  $\sigma$ -algebra  $\mathcal{F}$  is generated by W.

For any integer  $q \in \mathbb{N} \cup \{0\}$ , we denote by  $\mathcal{H}_q$  the *q*th *Wiener* chaos of *W*.  $\mathcal{H}_q$  is the closed linear subspace of  $L^2(\Omega)$  generated by the family of random variables  $\{H_q(W(h)), h \in \mathfrak{H}, \|h\|_{\mathfrak{H}} = 1\}$ , with  $H_q$  the *q*-th Hermite polynomial given by

$$H_q(x) = (-1)^q e^{\frac{x^2}{2}} \frac{d^q}{dx^q} \left( e^{-\frac{x^2}{2}} \right).$$

Let  $\mathcal{S}$  be the set of all cylindrical random variables of the form

$$F = g(W(h_1), \ldots, W(h_n)),$$

where  $n \ge 1$ ,  $h_i \in \mathfrak{H}$ , and g is infinitely differentiable such that all its partial derivatives have polynomial growth. The Malliavin derivative of F is the element of  $L^2(\Omega; \mathfrak{H})$  defined by

$$DF = \sum_{i=1}^{n} \frac{\partial g}{\partial x_i} (W(h_1), \dots, W(h_n)) h_i.$$

By iteration, for every  $m \ge 2$ , we define the *m*th derivative  $D^m F$ . This is an element of  $L^2(\Omega; \mathfrak{H}^{\odot m})$ , where  $\mathfrak{H}^{\odot m}$  designates the symmetric *m*th tensor product of  $\mathfrak{H}$ . For  $m \ge 1$  and  $p \ge 1$ ,  $\mathbb{D}^{m,p}$  denote the closure of S with respect to the norm  $\|\cdot\|_{m,p}$  defined by

$$\|F\|_{m,p}^{p} = \mathbf{E}[|F|^{p}] + \sum_{j=1}^{m} \mathbf{E}\left[\|D^{j}F\|_{\mathfrak{H}^{\otimes j}}^{p}\right]$$

Set  $\mathbb{D}^{\infty} = \bigcap_{m,p} \mathbb{D}^{m,p}$ . One can then extend the definition of  $D^m$  to  $\mathbb{D}^{m,p}$ . When m = 1, one simply write D instead of  $D^1$ . As a consequence of the hypercontractivity property of the Ornstein-Uhlenbeck semigroup, all the  $\|\cdot\|_{m,p}$ -norms are equivalent in any *finite* sum of Wiener chaoses.

Finally, let us recall that the Malliavin derivative D satisfies the following *chain rule*: if  $\varphi : \mathbb{R}^n \to \mathbb{R}$  is in  $\mathcal{C}_b^1$  (that is, belongs to the set of continuously differentiable functions with a bounded derivative) and if  $\{F_i\}_{i=1,...,n}$  is a vector of elements of  $\mathbb{D}^{1,2}$ , then  $\varphi(F_1,\ldots,F_n) \in \mathbb{D}^{1,2}$  and

$$D\varphi(F_1,\ldots,F_n) = \sum_{i=1}^n \frac{\partial \varphi}{\partial x_i}(F_1,\ldots,F_n) DF_i$$

Let  $\{F_n\}$  be a sequence of random variables belonging to a fixed chaos of order greater than or equal to 2.

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Let  $\{F_n\}$  be a sequence of random variables belonging to a fixed chaos of order greater than or equal to 2. Suppose  $\mathbb{E}[F_n^2] = 1$  and  $\lim_{n\to\infty} \mathbb{E}[F_n^4] = 3$ .

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$$\sup_{n}\mathbb{E}\left[\|DF_{n}\|^{-4-\epsilon}\right]<\infty.$$

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$$\sup_{n} \mathbb{E}\left[\|DF_{n}\|^{-4-\epsilon}\right] < \infty.$$

Then, there exists a constant c such that for all  $n \ge 1$ ,

$$\sup_{x\in\mathbb{R}}|p_{F_n}(x)-\phi(x)|\leq c\sqrt{\mathbb{E}[F_n^4]-3}.$$

(ii) Suppose that for all  $p \ge 1$ ,

$$\sup_{n} \mathbb{E}\left[\|DF_{n}\|^{-p}\right] < \infty.$$

Then, for any  $m \ge 0$ , there exists a constant  $c_m$  such that for all  $n \ge 1$ ,

$$\sup_{x\in\mathbb{R}}|p_{F_n}^{(m)}(x)-\phi^{(m)}(x)|\leq c_m\sqrt{\mathbb{E}[F_n^4]}-3.$$

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Hu, Y.; Lu, F. and Nualart, D.

Convergence of densities of some functionals of Gaussian processes.

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J. Funct. Anal. 266 (2014), 814-875.

# A key lemma

Our future computations will heavily rely on an efficient way to compute conditional expectations. Towards this aim, we state here some general results. Let us start with a decomposition for Hermite polynomials:

#### Lemma

For any  $q \ge 1$ , let  $H_q$  be the Hermite polynomial. Consider  $y, z \in \mathbb{R}$  and two real parameters a, b with  $a^2 + b^2 = 1$ . Then the following relation holds true:

$$H_q(ay+bz) = \sum_{\ell=0}^q \binom{q}{\ell} a^{q-\ell} b^\ell H_{q-\ell}(y) H_\ell(z).$$

## short proof

By the definition of the Hermite polynomials, we have

$$e^{aty-\frac{(at)^2}{2}} = \sum_{i=0}^{\infty} (at)^i H_i(y), \text{ and } e^{tbz-\frac{(bt)^2}{2}} = \sum_{j=0}^{\infty} (bt)^j H_j(z);$$
  
 $e^{t(ay+bz)-t^2/2} = \sum_{q=0}^{\infty} t^q H_q(ay+bz).$ 

Since  $a^2 + b^2 = 1$ , we obviously have  $e^{aty - \frac{(at)^2}{2}}e^{tbz - \frac{(bt)^2}{2}} = e^{t(ay+bz)-t^2/2}$ . Thus, we have

$$\sum_{q=0}^{\infty}t^{q}H_{q}(ay+bz)=\sum_{i=0}^{\infty}(at)^{i}H_{i}(y)\sum_{j=0}^{\infty}(bt)^{j}H_{j}(z),$$

which easily yields the desired identity.

This is to be used in the following computation of conditional expectations:

#### Proposition

Let Y and Z be two centered Gaussian random variables such that Y is measurable with respect to a  $\sigma$ -algebra  $\mathcal{G} \subset \mathcal{F}$  and Z is independent of  $\mathcal{G}$ . Assume that  $\mathbf{E}[Y^2] = \mathbf{E}[Z^2] = 1$ . Then for any  $q \geq 1$ , and real parameters a, b such that  $a^2 + b^2 = 1$ , we have:

$$\mathbf{E}[H_q(aY+bZ)|\mathcal{G}]=a^qH_q(Y).$$

# Short proof

Apply the key lemma in order to decompose  $H_q(aY + bZ)$ . Then identity follows easily from the fact that Y is  $\mathcal{G}$ -measurable, Z is independent from  $\mathcal{G}$  and Hermite polynomials have 0 mean under a centered Gaussian measure except for  $H_0 \equiv 1$ .

# Carbery-Wright inequality

#### Proposition

Let  $X = (X_1, \ldots, X_n)$  be a Gaussian random vector in  $\mathbb{R}^n$  and  $Q : \mathbb{R}^n \to \mathbb{R}$  a polynomial of degree at most m. Then there is a universal constant c > 0 such that:

$$(\mathbf{E}[|Q(X_1,...,X_n)|])^{\frac{1}{m}} \mathbf{P}(|Q(X_1,...,X_n)| \le x) \le c m x^{\frac{1}{m}}$$

for all x > 0.

# Sketch of the proof

Step 1: Computation of the Malliavin norm.

$$DV_n^{d,q} = \frac{1}{\sqrt{n}} \sum_{k=0}^{n-1} f'(X_k) \left( \sum_{j \ge 0} \psi_j \varepsilon^{k-j} \right)$$
$$= \frac{1}{\sqrt{n}} \sum_{l=-\infty}^{n-1} \left( \sum_{k=l^+}^{n-1} \psi_{k-l} f'(X_k) \right) \varepsilon^l,$$

where  $I^+ = \max\{I, 0\}$ .

$$\left\langle DV_n^{d,q}, DV_n^{d,q} \right\rangle_{\mathfrak{H}} = \frac{1}{n} \sum_{k_1,k_2=0}^{n-1} f'(X_{k_1}) \rho(k_1-k_2) f'(X_{k_2}),$$

$$\left\langle DV_n^{d,q}, DV_n^{d,q} \right\rangle_{\mathfrak{H}} = \frac{1}{n} \sum_{\ell=-\infty}^{n-1} \left( \sum_{k=\ell^+}^{n-1} \psi_{k-\ell} f'(X_k) \right)^2.$$

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Rearranging terms (namely, change  $k - \ell$  to k and then  $n - \ell - 1$  to m), we end up with:

$$\begin{split} \left\langle DV_{n}^{d,q}, DV_{n}^{d,q} \right\rangle_{\mathfrak{H}} &\geq \frac{1}{n} \sum_{\ell=0}^{n-1} \left( \sum_{k=0}^{n-\ell-1} f'(X_{\ell+k}) \psi_{k} \right)^{2} \\ &= \frac{1}{n} \sum_{m=0}^{n-1} \left( \sum_{k=0}^{m} f'(X_{n-1-(m-k)}) \psi_{k} \right)^{2} \equiv A_{n}. \end{split}$$

As a last preliminary step we resort to the fact that  $X = \{X_k; k \in \mathbb{N} \cup \{0\}\}$  is a Gaussian stationary sequence, which allows to assert that  $A_n$  is identical in law to  $B_n$  with

$$B_n := \frac{1}{n} \sum_{m=0}^{n-1} \left( \sum_{k=0}^m f'(X_{m-k}) \psi_k \right)^2 = \frac{1}{n} \sum_{m=0}^{n-1} \left( \sum_{k=0}^m f'(X_k) \psi_{m-k} \right)^2$$

We will now bound the negative moments of  $B_n$ .

#### Step 2: Block decomposition.

Fix thus an integer  $N \ge 1$  and let M = [n/N] be the integer part of n/N. Then  $n \ge NM$  and as a consequence,

$$B_{n} = \frac{1}{n} \sum_{m=0}^{n-1} \left( \sum_{k=0}^{m} f'(X_{k}) \psi_{m-k} \right)^{2}$$
  

$$\geq \frac{1}{n} \sum_{i=0}^{N-1} \sum_{m=iM}^{(i+1)M-1} \left( \sum_{k=0}^{m} f'(X_{k}) \psi_{m-k} \right)^{2}.$$

For  $i = 0, \ldots, N - 1$ , define

$$B_n^i = \frac{1}{n} \sum_{m=iM}^{(i+1)M-1} \left( \sum_{k=0}^m f'(X_k) \psi_{m-k} \right)^2$$

so that 
$$B_n \geq \sum_{i=0}^{N-1} B_n^i$$
.

Then it is readily checked that:

$$(B_n)^{-\frac{p}{2}} \leq \prod_{i=0}^{N-1} (B_n^i)^{-\frac{p}{2N}}$$

we obtain:

$$\mathbf{E}\left[(B_n)^{-\frac{p}{2}}\right] \leq \mathbf{E}\left[\prod_{i=0}^{N-1} (B_n^i)^{-\frac{p}{2N}}\right]$$
$$= \mathbf{E}\left[\mathbf{E}\left[(B_n^{N-1})^{-\frac{p}{2N}} | \mathcal{F}_{(N-1)M}\right] \prod_{i=0}^{N-2} (B_n^i)^{-\frac{p}{2N}}\right]$$

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Step 3: Application of Carbery-Wright. Let us go back to the particular situation of  $f = \sum_{j=d}^{q} a_j H_j$ , which means in particular that  $f' = \sum_{j=d}^{q} j a_j H_{j-1}$ . First, we notice

$$\mathbf{E} \left[ (B_n^{N-1})^{-\frac{p}{2N}} | \mathcal{F}_{(N-1)M} \right]$$
  
  $\leq 1 + \frac{p}{2N} \int_0^1 \mathbf{P} \left( B_n^{N-1} \leq x | \mathcal{F}_{(N-1)M} \right) x^{-\frac{p}{2N} - 1} dx .$ 

Since  $B_n^{N-1}$  is a polynomial of order m = 2(q-1), Carbery-Wright's inequality yields:

$$\mathbf{P}\left(B_{n}^{N-1} \leq x | \mathcal{F}_{(N-1)M}\right) \leq \frac{c \, x^{\frac{1}{2(q-1)}}}{\left[\mathbf{E}\left(B_{n}^{N-1} | \mathcal{F}_{(N-1)M}\right)\right]^{\frac{1}{2(q-1)}}}$$

Step 4: Estimates for the conditional expectation. We now estimate the conditional expectation  $\mathbf{E}[B_n^{N-1}|\mathcal{F}_{(N-1)M}]$ . We have:

$$\begin{split} \mathbf{E} & \left[ B_n^{N-1} | \mathcal{F}_{(N-1)M} \right] \\ &= \frac{1}{n} \sum_{m=(N-1)M}^{NM-1} \mathbf{E} \left[ \left( \sum_{k=0}^m f'(X_k) \psi_{m-k} \right)^2 \Big| \mathcal{F}_{(N-1)M} \right] \\ &\geq \frac{1}{n} \sum_{m=(N-1)M}^{NM-1} A_m, \end{split}$$

where we have set

$$A_m = \operatorname{Var}\left(\sum_{k=(N-1)M}^m f'(X_k)\psi_{m-k}\Big|\mathcal{F}_{(N-1)M}\right).$$

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Furthermore, notice that

$$f'(X_k) = f'\left(\sum_{\ell=-\infty}^k \psi_{k-i} w_i\right) = f'(Y_k + Z_k),$$

where  $Y_k = \sum_{i=-\infty}^{(N-1)M-1} \psi_{k-i} w_i$  is  $\mathcal{F}_{(N-1)M}$ -measurable and  $Z_k = \sum_{i=(N-1)M}^k \psi_{k-i} w_i$  is independent of  $\mathcal{F}_{(N-1)M}$ . Recalling that  $f' = \sum_{i=d}^q j a_i H_{j-1}$ . This gives:

$$\begin{aligned} & \mathcal{H}_{q-1}(X_k) - \mathbf{E}[\mathcal{H}_{q-1}(X_k) | \mathcal{F}_{(N-1)M}] \\ &= \sum_{j=d}^{q} \sum_{\ell=1}^{j-1} j^2 a_j \binom{j-1}{\ell} \sigma_{Y_k}^{j-1-\ell} \mathcal{H}_{j-1-\ell}(\widetilde{Y}_k) \sigma_{Z_k}^{\ell} \mathcal{H}_{\ell}(\widetilde{Z}_k), \end{aligned}$$

where  $\sigma_{Y_k} = [Var(Y_k)]^{1/2}$ ,  $\sigma_{Z_k} = [Var(Z_k)]^{1/2}$ ,  $\widetilde{Y}_k = Y_k / \sigma_{Y_k}$  and  $\widetilde{Z}_k = Z_k / \sigma_{Z_k}$ .

Therefore,

$$A_{m} = \mathbf{E} \left[ \left( \sum_{k=(N-1)M}^{m} \sum_{j=d}^{q} \sum_{\ell=1}^{j-1} a_{j,\ell,k} H_{j-1-\ell}(\widetilde{Y}_{k}) H_{\ell}(\widetilde{Z}_{k}) \psi_{m-k} \right)^{2} | \mathcal{F}_{(N-1)M} \right]$$
$$= \mathbf{E} \left[ \left( \sum_{\ell=1}^{q-1} \sum_{k=(N-1)M}^{m} \sum_{j=(\ell+1)\vee d}^{q} a_{j,\ell,k} H_{j-1-\ell}(\widetilde{Y}_{k}) H_{\ell}(\widetilde{Z}_{k}) \psi_{m-k} \right)^{2} | \mathcal{F}_{(N-1)M} \right]$$

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where we have set  $a_{j,\ell,k} = j^2 a_j {j-1 \choose \ell} \sigma_{Y_k}^{j-1-\ell} \sigma_{Z_k}^{\ell}$ .

Recall that the random variables  $\widetilde{Y}_k$  are  $\mathcal{F}_{(N-1)M}$ -measurable while the random variables  $\widetilde{Z}_k$  are independent of  $\mathcal{F}_{(N-1)M}$ . By decorrelation properties of Hermite polynomials we thus get:

$$A_{m} = \sum_{\ell=1}^{q-1} \mathbf{E} \left[ \left( \sum_{k=(N-1)M}^{m} \sum_{j=(\ell+1)\vee d}^{q} a_{j,\ell,k} H_{j-1-\ell}(\widetilde{Y}_{k}) H_{\ell}(\widetilde{Z}_{k}) \psi_{m-k} \right)^{2} \Big| \mathcal{F}_{(N-1)M} \right]$$

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and we trivially lower bound this quantity by taking the term corresponding to  $\ell = q - 1$ . In this situation the sum  $\sum_{j=(\ell+1)\vee d}^{q}$  is reduced to the term corresponding to j = q, and since  $a_{q,q-1,k} = q^2 a_q \sigma_{Z_k}^{q-1}$  we obtain:

$$A_m \ge \mathbf{E} \left[ \left( \sum_{k=(N-1)M}^m q^2 a_q \, \sigma_{Z_k}^{q-1} \, H_{q-1}(\widetilde{Z}_k) \, \psi_{m-k} \right)^2 \Big| \mathcal{F}_{(N-1)M} \right]$$
$$= q^4 \, a_q^2 \, \mathbf{E} \left[ \left( \sum_{k=(N-1)M}^m \sigma_{Z_k}^{q-1} \, H_{q-1}(\widetilde{Z}_k) \, \psi_{m-k} \right)^2 \right].$$

We now invoke the identity  $\mathbf{E}[H_p(\tilde{Z}_{k_1})H_p(\tilde{Z}_{k_2})] = \frac{1}{p!}(\mathbf{E}[\tilde{Z}_{k_1}\tilde{Z}_{k_2}])^p$ in order to obtain

$$A_m \geq \frac{q^5 a_q^2}{q!} \sum_{k_1, k_2 = (N-1)M}^m \sigma_{Z_{k_1}}^{q-1} \sigma_{Z_{k_2}}^{q-1} \mathbf{E} \left[ \widetilde{Z}_{k_1} \widetilde{Z}_{k_2} \right]^{q-1} \psi_{m-k_1} \psi_{m-k_2}.$$

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Furthermore, it is readily checked that:

$$\mathbf{E}\left[\widetilde{Z}_{k_1}\,\widetilde{Z}_{k_2}\right] = \frac{1}{\sigma_{Z_{k_1}}\sigma_{Z_{k_2}}}\sum_{i=(N-1)M}^{k_1\wedge k_2}\psi_{k_1-i}\,\psi_{k_2-i},$$

#### and thus

$$\begin{split} A_m &\geq \frac{q^5 \, a_q^2}{q!} \sum_{k_1, k_2 = (N-1)M}^m \left( \sum_{i=(N-1)M}^{k_1 \wedge k_2} \psi_{k_1 - i} \, \psi_{k_2 - i} \right)^{q-1} \psi_{m-k_1} \psi_{m-k_2} \\ &= \frac{q^5 \, a_q^2}{q!} \sum_{i_1, \dots, i_{q-1} = (N-1)M}^m \sum_{k_1, k_2 = \max(i_1, \dots, i_{q-1})}^m \psi_{m-k_1} \psi_{m-k_2} \prod_{j=1}^{q-1} \psi_{k_1 - i_j} \, \psi_{k_2 - i_j} \\ &= \frac{q^5 \, a_q^2}{q!} \sum_{i_1, \dots, i_{q-1} = (N-1)M}^m \left( \sum_{k=\max(i_1, \dots, i_{q-1})}^m \psi_{m-k} \prod_{j=1}^{q-1} \psi_{k-i_j} \right)^2. \end{split}$$

Here again, this sum of squares is trivially lower bounded by taking the term corresponding to  $i_1 = \cdots = i_{q-1} = m$ , which yields:

$$A_m \geq c_{a,q,\psi} \quad ext{with} \quad c_{a,q,\psi} \equiv rac{q^5 \, a_q^2}{q!} \, \psi_0^{2q} > 0.$$

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Step 5: Conclusion. Recalling that N is a given integer whose exact value will be fixed below, we get:

$$\mathsf{E}\left[B_n^{N-1}|\mathcal{F}_{(N-1)M}\right] \geq \frac{M\,c_{\mathsf{a},q,\psi}}{n} \geq c_{\mathsf{a},q,\psi,N} > 0,$$

as long as N stays bounded. We then get:

$$\mathbf{P}\left(B_{n}^{N-1} \leq x | \mathcal{F}_{(N-1)M}\right)$$

$$\leq 1 + \frac{p c_{a,q,\psi,N}}{2N} \int_{0}^{1} x^{\frac{1}{2(q-1)} - \frac{p}{2N} - 1} dx = c_{a,q,\psi,N,p} < \infty,$$

where we have chosen N such that  $\frac{p}{2N} < \frac{1}{2(q-1)}$ . Iterating this bound, we have thus obtained:

$$\mathsf{E}\left[(B_n)^{-\frac{p}{2}}\right] \leq c_{\boldsymbol{a},\boldsymbol{q},\boldsymbol{\psi},\boldsymbol{N},\boldsymbol{p}}^{\boldsymbol{N}},$$

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which is a finite quantity.

Finally recall from Step 1 that  $\mathbf{E}[(B_n)^{-\frac{p}{2}}] = \mathbf{E}[\|DV_n^{d,q}\|_{\mathfrak{H}}^{-p}]$ , which finishes the proof.

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# **THANKS**

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