



The deviation matrix, Poisson's equation, and Quasi-birth-death processes (QBDs)

Yuanyuan Liu

Institute of Probability and Statistics, School of Mathematics and Statistics, Central South University

the Ninth Workshop on Markov Processes and Related Topics
SWJTU and BNU, July 6-13, 2013

Joint work with Sarah Dendievel and Guy Latouche



中南大學
CENTRAL SOUTH UNIVERSITY



The deviation matrix, Poisson's equation, and Quasi-birth-death processes (QBDs)

Yuanyuan Liu

Institute of Probability and Statistics, School of Mathematics and Statistics, Central South University

the Ninth Workshop on Markov Processes and Related Topics
SWJTU and BNU, July 6-13, 2013

Joint work with Sarah Dendievel and Guy Latouche

Poisson's equation

Let $\{\Phi_0, \Phi_1, \Phi_2, \dots\}$ be a discrete-time Markov chain.

Poisson's equation:

$$(I - P)\underline{x} = \underline{g}$$

where

- P is the transition matrix $P \geq 0$, $P\underline{1} = \underline{1}$,
- \underline{g} is a given vector indexed by the state space

Assume

- denumerable state space, irreducible, **positive recurrent**
- $\underline{\pi}^t \underline{g} = 0$. w.l.g.

Link with Central limit theorem

Define $S_n = \frac{1}{n} \sum_{0 \leq t \leq n-1} \mathbf{g} \Phi_t$

Strong Law of Large Number:

$$S_n \rightarrow \underline{\pi}^t \underline{g} \quad \text{a.s. for } n \rightarrow \infty$$

Central limit theorem:

$$\sqrt{n}(S_n - \underline{\pi}^t \underline{g}) \Rightarrow N(0, \sigma_{\underline{g}}^2) \quad \text{for } n \rightarrow \infty$$

where $\sigma_{\underline{g}}^2 = \sum_i \pi_i (2h_i \bar{g}_i - \bar{g}_i^2)$

$$(I - P)\underline{h} = \underline{\bar{g}} \quad \underline{\bar{g}} = \underline{g} - (\underline{\pi}^t \underline{g})\underline{1}$$



Link with perturbation

We need P is aperiodic.

Take perturbation E : matrix s.t. $Q = P + E$ is stochastic, irreducible, positive recurrent, ...

Define $\underline{\alpha}$ to be the stationary probability vector of Q

If E is sufficiently small,

$$\underline{\alpha}^t = \underline{\pi}^t \sum_{n \geq 0} (E\mathcal{D})^n$$

where $\mathcal{D} := \sum_{n \geq 0} (P^n - \underline{1}\underline{\pi}^t)$ is the deviation matrix such that

$$(I - P)\mathcal{D} = I - \underline{1}\underline{\pi}^t \quad \underline{\pi}^t \mathcal{D} = \underline{0}$$



Outline

- 1 Markov chains: Finite state space
- 2 Markov chains: Infinite state space
- 3 Properties of QBDs
- 4 Solving Poisson's equation for QBDs
- 5 Illustration



Finite state space is easy

Poisson's equation

$$(I - P)\underline{x} = \underline{g}$$

P stochastic, irreducible, finite size: \underline{x} is unique, up to an additive constant

$$\underline{x} = (I - P)^\# \underline{g} + c \underline{1}$$

- $(I - P)^\#$ is the group inverse of $(I - P)$
- c is an arbitrary constant
- actually, $\underline{\pi}^\dagger \underline{g} = 0$ otherwise system doesn't make sense

Groupe inverses and deviation matrix

Groupe inverse of A (unique when it exists):

$$(1) AA^\#A = A, \quad A^\#AA^\# = A^\#$$

$$(2) AA^\# = A^\#A$$

- For properties and computation:

Campbell and Meyer, *Generalized Inverses of Linear Transformations*, 1979

- Irreducible finite MC: $(I - P)^\#$ exists and is unique solution to

$$(I - P)(I - P)^\# = I - \underline{1}\pi^t, \quad \pi^t(I - P)^\# = \underline{0}$$

Deviation matrix: (in addition, P is **non-periodic**)

$$\mathcal{D} := \sum_{n \geq 0} (P^n - \underline{1}\pi^t) = (I - P)^\#.$$

Constructive solution

GLynn and Meyn (1996): Assume $\underline{\pi}^{\dagger}|\underline{g}| < \infty$

(i) Take j to be an arbitrary state and T to be its first return time
 One solution of the Poisson equation $(I - P)\underline{x} = \underline{g}$ is given by

$$x_i = \mathbb{E}\left[\sum_{0 \leq n < T} g_{\Phi_n} | \Phi_0 = i\right]$$

$$x_j = 0.$$

(ii) (uniqueness:) solutions are given up to an arbitrary constant.

Comments of (i):

It is not convenient to consider single state j for matrix-analytic models.

Censoring — a.k.a. Schur complementation

Take subset of states A , T first return time to A ,

$$P = \begin{bmatrix} P_{AA} & P_{AB} \\ P_{BA} & P_{BB} \end{bmatrix} \quad N_B = \sum_{n \geq 0} P_{BB}^n$$

$$\beta_i = \mathbb{E} \left[\sum_{0 \leq n < T} g_{\Phi_n} \mid \Phi_0 = i \right]$$

Dendievel, Latouche and Liu (2013):

Theorem 1: One solution is given by

$$\underline{x}_A = \underline{\beta}_A + (P_{AA} + P_{AB} N_B P_{BA}) \underline{x}_A$$

$$\underline{x}_B = \underline{\beta}_B + N_B P_{BA} \underline{x}_A$$

Schur complementation, same as censoring for stationary distribution.

Deviation matrix

$$\mathcal{D} = \sum_{n \geq 0} (P^n - \underline{1}\underline{\pi}^t)$$

The deviation matrix exists (i.e. the series converges) if and only if $E[T^2(j)|\Phi_0 = j] < \infty$

Bhulai and Spieksma (2003)

Assume P is geometrically ergodic.

(i) \mathcal{D} is the unique solution of

$$(I - P)\mathcal{D} = I - \underline{1}\underline{\pi}^t, \quad \underline{\pi}^t\mathcal{D} = \underline{0}$$

(ii) A solution of $(I - P)\underline{x} = \underline{g}$, with $\underline{\pi}^t\underline{g} = 0$, $\underline{\pi}^t|\underline{g}| < \infty$, is $\underline{x} = \mathcal{D}\underline{g} + c\underline{1}$, where c is arbitrary.

In this way, \mathcal{D} is just like $(I - P)^\#$.



QBDs

QBDs are Markov chains on a two-dimensional state space

$$(n, \varphi) : \quad n = 0, 1, 2, \dots; \quad \varphi = 1, 2, \dots, M$$

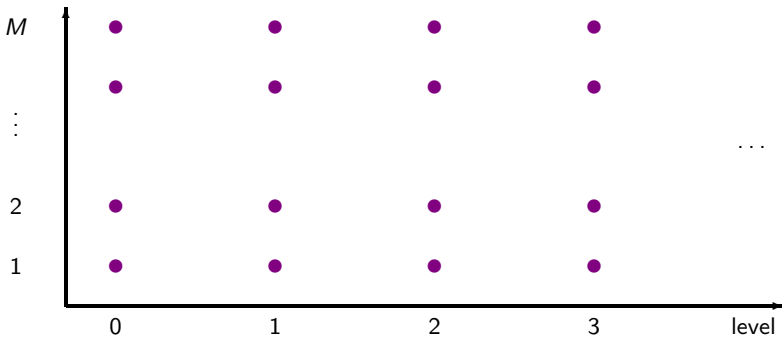
here $M < \infty$.

Often,

- n is length of a queue, named the **level**,
- φ may be many different things, named the **phase**.

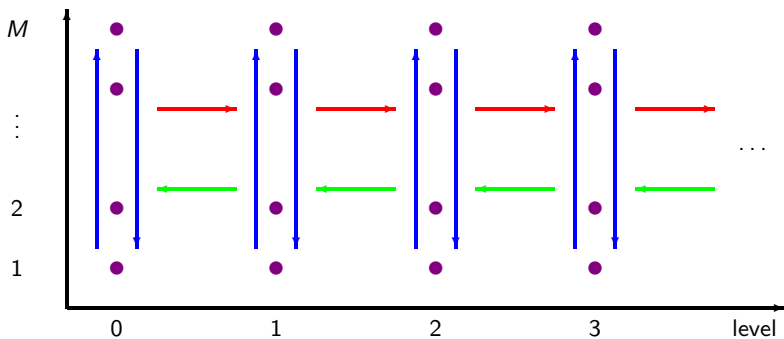


Transition graph (such as it is)





Transition graph (such as it is)



Transition matrix

Block-structured transition matrix:

$$P = \begin{bmatrix} A_* & A_1 & 0 & 0 & \cdots \\ A_{-1} & A_0 & A_1 & 0 & \\ 0 & A_{-1} & A_0 & A_1 & \ddots \\ 0 & 0 & A_{-1} & A_0 & \ddots \\ \vdots & & \ddots & \ddots & \ddots \end{bmatrix}$$

Transition probabilities:

$(A_1)_{ij}$ probability to go up from (n, i) to $(n + 1, j)$



Transition matrix

Block-structured transition matrix:

$$P = \begin{bmatrix} A_* & A_1 & 0 & 0 & \cdots \\ A_{-1} & A_0 & A_1 & 0 & \\ 0 & A_{-1} & A_0 & A_1 & \ddots \\ 0 & 0 & A_{-1} & A_0 & \ddots \\ \vdots & & \ddots & \ddots & \ddots \end{bmatrix}$$

Transition probabilities:

$(A_{-1})_{ij}$ probability to go down from (n, i) to $(n-1, j)$



Transition matrix

Block-structured transition matrix:

$$P = \begin{bmatrix} A_* & A_1 & 0 & 0 & \cdots \\ A_{-1} & A_0 & A_1 & 0 & \\ 0 & A_{-1} & A_0 & A_1 & \ddots \\ 0 & 0 & A_{-1} & A_0 & \ddots \\ \vdots & & \ddots & \ddots & \ddots \end{bmatrix}$$

Transition probabilities:

$(A_0)_{ij}$ probability to stay in level n , (n, i) to (n, j) , $n \neq 0$



Transition matrix

Block-structured transition matrix:

$$P = \begin{bmatrix} A_* & A_1 & 0 & 0 & \cdots \\ A_{-1} & A_0 & A_1 & 0 & \\ 0 & A_{-1} & A_0 & A_1 & \ddots \\ 0 & 0 & A_{-1} & A_0 & \ddots \\ \vdots & & \ddots & \ddots & \ddots \end{bmatrix}$$

Transition probabilities:

$(A_*)_{ij}$ probability to remain in level 0, $(0, i)$ to $(0, j)$

matrices for QBDs

Analysis makes extensive use of matrices

$$G_{ij} = P[T < \infty, \Phi_T = (0, j) | \Phi_0 = (1, i)],$$

$$R_{ij} = E\left[\sum_{0 \leq t < T} \mathbb{1}[\Phi_t = (1, j)] | \Phi_0 = (0, i)\right],$$

$$U = A_0 + A_1 G$$

T is the first return time to level 0.

Computing matrices G and R .

Latouche-Ramaswami's Algorithm (1993) is powerful.



QBDs

Our focus:

Look for “constructive” solution to

$$(I - P)\underline{x} = \underline{g},$$

where P is the transition matrix of an irreducible, aperiodic and positive recurrent QBD.



Solution 1: through the first return time

Let $A = \ell(0)$ denote the level 0 and $B = \mathbb{E} \setminus A$. Using Theorem 1, we have:

Dendievel, Latouche and Liu (2013)

Theorem 2. A solution of the Poisson's equation for a QBD is given by

$$\mathbf{h}_0 = (I - P^*)^\# [\mathbf{y}_0 - (\boldsymbol{\pi}\mathbf{g})\boldsymbol{\tau}_0] + \mathbf{1},$$

$$\mathbf{h}_n = \mathbf{y}_n - (\boldsymbol{\pi}\mathbf{g})\boldsymbol{\tau}_n + G^n \mathbf{h}_0, \quad n \geq 1,$$

where $P^* = B + A_1 G$, and $\mathbf{y}_n, \boldsymbol{\tau}_n$ are obtained by

$$\boldsymbol{\tau}_0 = \mathbf{1} + A_1 \boldsymbol{\tau}_1,$$

$$\boldsymbol{\tau}_n = \left[(I - G^n)(I - G)^\# + n\mathbf{1}\boldsymbol{\nu}^\top \right] (I - U)^{-1} (I - R)^{-1} \mathbf{1}, \quad n \geq 1,$$

$$\mathbf{y}_0 = \mathbf{g}_0 + A_1 \mathbf{y}_1,$$

$$\mathbf{y}_n = \sum_{i=0}^{n-1} G^i (I - U)^{-1} \left[\sum_{l \geq 0} R^l \mathbf{g}_{n-i+l} \right], \quad n \geq 1,$$

Solution 2: through the deviation matrix

Let $C = I - \underline{1}\pi^t$. Look for the deviation matrix by solving

$$(I - P)X = C$$

i.e. $X = \mathcal{D}$.

Details are a bit messy, focus on the structure.

$$\left[\begin{array}{c|ccc} I - A_* & -A_1 & 0 & \dots \\ \hline -A_{-1} & I - A_0 & -A_1 & \\ 0 & -A_{-1} & I - A_0 & \\ \vdots & & & \ddots \end{array} \right] \begin{bmatrix} X_0 \\ X_1 \\ X_2 \\ \vdots \end{bmatrix} = \begin{bmatrix} C_0 \\ C_1 \\ C_2 \\ \vdots \end{bmatrix}$$



Solution 2: through the deviation matrix

Let $C = I - \underline{1}\pi^t$. Look for the deviation matrix by solving

$$(I - P)X = C$$

i.e. $X = \mathcal{D}$.

Details are a bit messy, focus on the structure.

$$\left[\begin{array}{c|ccc} I - A_* & -A_1 & 0 & \dots \\ \hline -A_{-1} & I - A_0 & -A_1 & \\ 0 & -A_{-1} & I - A_0 & \\ \vdots & & & \ddots \end{array} \right] \begin{bmatrix} X_0 \\ X_1 \\ X_2 \\ \vdots \end{bmatrix} = \begin{bmatrix} C_0 \\ C_1 \\ C_2 \\ \vdots \end{bmatrix}$$

Schur complementation

$$\left[\begin{array}{c|ccc} I - A_* & -A_1 & 0 & \dots \\ \hline -A_{-1} & I - A_0 & -A_1 & \\ 0 & -A_{-1} & I - A_0 & \\ \vdots & & & \ddots \end{array} \right] \begin{bmatrix} X_0 \\ X_1 \\ X_2 \\ \vdots \end{bmatrix} = \begin{bmatrix} C_0 \\ C_1 \\ C_2 \\ \vdots \end{bmatrix}$$

Gaussian elimination:

$$\left[\begin{array}{cc} I - A_0 & -A_1 \\ -A_{-1} & I - A_0 \\ & \ddots \end{array} \right] \begin{bmatrix} X_1 \\ X_2 \\ \vdots \end{bmatrix} = \begin{bmatrix} C_1 \\ C_2 \\ \vdots \end{bmatrix} + \begin{bmatrix} A_{-1} \\ 0 \\ \vdots \end{bmatrix} X_0$$

Isolate $[X_1 \ X_2 \ \dots]$ and inject in first equation

Schur complementation (contd)

$$\begin{bmatrix} I - A_0 & -A_1 & & \\ -A_{-1} & I - A_0 & & \\ & & \ddots & \\ & & & \ddots \end{bmatrix} \begin{bmatrix} X_1 \\ X_2 \\ \vdots \end{bmatrix} = \begin{bmatrix} C_1 \\ C_2 \\ \vdots \end{bmatrix} + \begin{bmatrix} A_{-1} \\ 0 \\ \vdots \end{bmatrix} X_0$$

Need W such that

$$W \begin{bmatrix} I - A_0 & -A_1 & & \\ -A_{-1} & I - A_0 & & \\ & & \ddots & \\ & & & \ddots \end{bmatrix} = I$$

Or $W = \sum_{n \geq 0} H^n$,

$$H = \begin{bmatrix} A_0 & A_1 & & \\ A_{-1} & A_0 & & \\ & & \ddots & \\ & & & \ddots \end{bmatrix}$$



W : the matrix of the expected sojourn times

$W_{(n,i)(k,j)}$ is expected number of visits to (k, j) before level zero, starting from (n, i) . A bit calculation yields

$$W_{nk} = \begin{cases} G^{n-k} W_{kk}, & n > k, \\ W_{nn} R^{k-n}, & n < k, \\ \sum_{\nu=0}^{n-1} G^{\nu} (I - U)^{-1} R^{\nu}, & n = k. \end{cases}$$

Probability explanation of the first case: $n > k$

$$n \xrightarrow{G} n-1 \xrightarrow{G} \cdots \xrightarrow{G} k$$

- (a) go down $n - k$ levels from n to k and
 (b) start counting

$$W_{nk} = G^{n-k} W_{kk}$$



W : the matrix of the expected sojourn times

$W_{(n,i)(k,j)}$ is expected number of visits to (k, j) before level zero, starting from (n, i) . A bit calculation yields

$$W_{nk} = \begin{cases} G^{n-k} W_{kk}, & n > k, \\ W_{nn} R^{k-n}, & n < k, \\ \sum_{\nu=0}^{n-1} G^{\nu} (I - U)^{-1} R^{\nu}, & n = k. \end{cases}$$

Probability explanation of the first case: $n > k$

$$n \xrightarrow{G} n-1 \xrightarrow{G} \cdots \xrightarrow{G} k$$

- (a) go down $n - k$ levels from n to k and
 (b) start counting

$$W_{nk} = G^{n-k} W_{kk}$$

Schur complementation (end)

$$\left[\begin{array}{c|ccc} I - A_* & -A_1 & 0 & \dots \\ \hline -A_{-1} & I - A_0 & -A_1 & \\ 0 & -A_{-1} & I - A_0 & \\ \vdots & & & \ddots \end{array} \right] \begin{bmatrix} X_0 \\ X_1 \\ X_2 \\ \vdots \end{bmatrix} = \begin{bmatrix} C_0 \\ C_1 \\ C_2 \\ \vdots \end{bmatrix}$$

First equation becomes

$$(I - P_*)X_0 = \sum_{i \geq 0} R^i C_i,$$

where $P_* = B + A_1 G$ is the transition matrix of the censored Markov chain.

The expression of the deviation matrix

For deviation matrix, let $C = I - \underline{1}\underline{\pi}^t$ (in $(I - P)X = C$).

Dendievel, Latouche and Liu (2013)

Theorem 3. The deviation matrix \mathcal{D} is given by $\mathcal{D} = (I - \underline{1}\underline{\pi}^t)K$, where

$$K_{0k} = (I - P_*)^\# (I - \underline{\tau}_0 \underline{\pi}_0^t) R^k \quad k \geq 0$$

$$K_{n0} = -\underline{\tau}_n \underline{\pi}_0^t + G^n K_{00} \quad n \geq 1$$

$$K_{nk} = W_{nk} - \underline{\tau}_n \underline{\pi}_k^t + G^n K_{0k} \quad n, k \geq 1,$$

- One solution of the Poisson equation is given by $\underline{x} = \mathcal{D}\underline{g} + c\underline{1}$.

The expression of the deviation matrix

For deviation matrix, let $C = I - \underline{1}\underline{\pi}^t$ (in $(I - P)X = C$).

Dendievel, Latouche and Liu (2013)

Theorem 3. The deviation matrix \mathcal{D} is given by $\mathcal{D} = (I - \underline{1}\underline{\pi}^t)K$, where

$$K_{0k} = (I - P_*)^\# (I - \underline{\tau}_0 \underline{\pi}_0^t) R^k \quad k \geq 0$$

$$K_{n0} = -\underline{\tau}_n \underline{\pi}_0^t + G^n K_{00} \quad n \geq 1$$

$$K_{nk} = W_{nk} - \underline{\tau}_n \underline{\pi}_k^t + G^n K_{0k} \quad n, k \geq 1,$$

- One solution of the Poisson equation is given by $\underline{x} = \mathcal{D}\underline{g} + c\underline{1}$.

The expression of the deviation matrix

For deviation matrix, let $C = I - \underline{1}\underline{\pi}^t$ (in $(I - P)X = C$).

Dendievel, Latouche and Liu (2013)

Theorem 3. The deviation matrix \mathcal{D} is given by $\mathcal{D} = (I - \underline{1}\underline{\pi}^t)K$, where

$$K_{0k} = (I - P_*)^\# (I - \underline{\tau}_0 \underline{\pi}_0^t) R^k \quad k \geq 0$$

$$K_{n0} = -\underline{\tau}_n \underline{\pi}_0^t + G^n K_{00} \quad n \geq 1$$

$$K_{nk} = W_{nk} - \underline{\tau}_n \underline{\pi}_k^t + G^n K_{0k} \quad n, k \geq 1,$$

- One solution of the Poisson equation is given by $\underline{x} = \mathcal{D}\underline{g} + c\underline{1}$.

The expression of the deviation matrix

For deviation matrix, let $C = I - \underline{1}\underline{\pi}^t$ (in $(I - P)X = C$).

Dendievel, Latouche and Liu (2013)

Theorem 3. The deviation matrix \mathcal{D} is given by $\mathcal{D} = (I - \underline{1}\underline{\pi}^t)K$, where

$$K_{0k} = (I - P_*)^\# (I - \underline{\tau}_0 \underline{\pi}_0^t) R^k \quad k \geq 0$$

$$K_{n0} = -\underline{\tau}_n \underline{\pi}_0^t + G^n K_{00} \quad n \geq 1$$

$$K_{nk} = W_{nk} - \underline{\tau}_n \underline{\pi}_k^t + G^n K_{0k} \quad n, k \geq 1,$$

- One solution of the Poisson equation is given by $\underline{x} = \mathcal{D}\underline{g} + c\underline{1}$.

Illustration

Assume QBD is a queueing system, **level = number of customers**.

Define

$$m_{\ell,j} = \frac{1}{L} \sum_{n \geq 0} (\mathbb{E}[Y_n | Y_0 = \ell, \varphi_0 = j] - L)$$

where L is stationary expected number of customers.

One has

$$\underline{m} = \frac{1}{L} \mathcal{D} \underline{\gamma} \quad \text{with } \underline{\gamma}_n = n \underline{1}.$$

Special form of $\underline{\gamma}$ allows for **further simplification** and makes it possible to compute \underline{m} with **finite computations**.



PH/M/1 queue

Example: PH/M/1 queue, services are exponential, interarrival times are PH($\underline{\tau}, T$) with

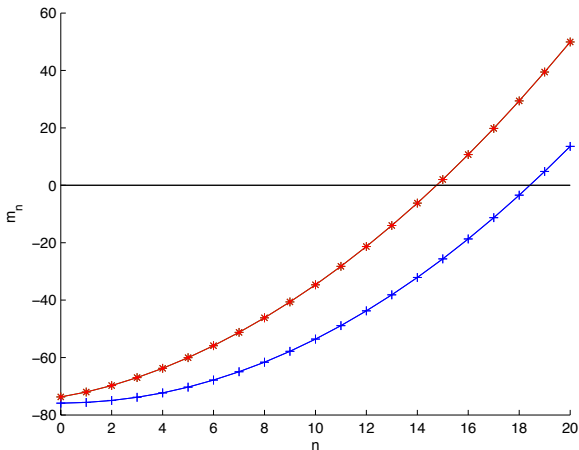
$$\underline{\tau} = [0.1127 \quad 0.8873],$$

$$T = \begin{bmatrix} -0.2254 & 0 \\ 0 & -1.7746 \end{bmatrix}.$$

Service rate = 1.2.

Traffic coefficient = 0.8333.

PH/M/1 queue (Contd)



$L = 11.1$. Plot of $m_{n,i}$: blue line for phase $i = 1$; red line for phase $i = 2$.

