

Branching structure for an $(L-1)$ random walk in random environment and its applications

Huaming Wang

(A joint work with Wenming Hong)

School of Mathematical Sciences, Beijing Normal
University & College of Business, Beijing Union
University

July 19, 2010



- ① Backgrounds and Motivations
- ② The construction of the branching structure for (L-1) RWRE
- ③ Applications of the branching structure of (L-1) RWRE (Stable law and LLN)
- ④ Tail estimates of MBPIRE



Backgrounds

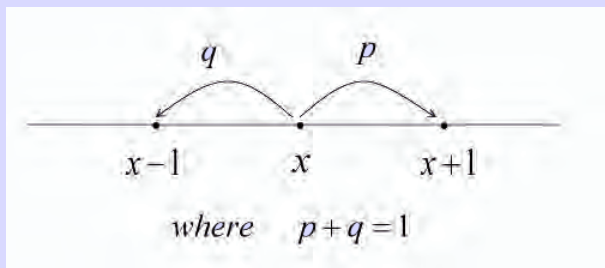
Nearest neighbor setting

(1) Simple random walk:

$$S_n = \sum_{i=1}^n \xi_i,$$

where ξ_i i.i.d., satisfying $P(\xi_i = 1) = p, P(\xi_i = -1) = q$.

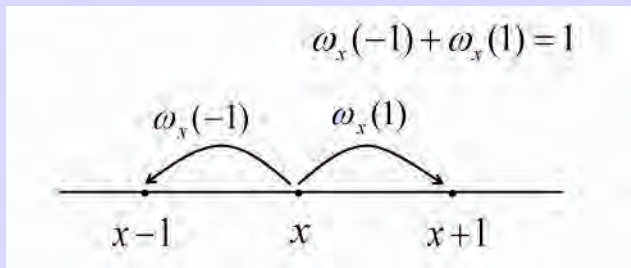
$$P(S_n = j | S_{n-1} = i) = \begin{cases} p & \text{if } j = i + 1, \\ q & \text{if } j = i - 1. \end{cases}$$



Simple RW with nonhomogeneous environment

(2) Simple RW with nonhomogeneous environment

$$\begin{aligned}P_{\omega}(X_{n+1} = x + 1 | X_n = x) &= \omega_x(1); \\P_{\omega}(X_{n+1} = x - 1 | X_n = x) &= \omega_x(-1).\end{aligned}\tag{1}$$



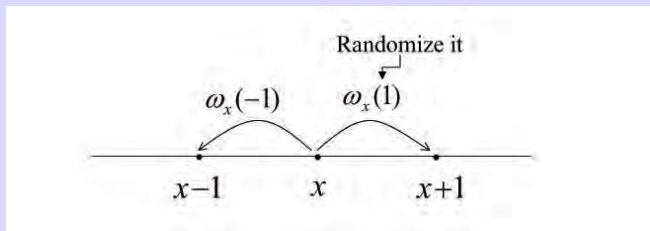
Here X_n cannot be written as sums of i.i.d. random variables any longer.



Random Walks in Random Environment

Randomizing the environment $\{\omega_x\}_{x \in \mathbb{Z}}$, letting $\{\omega_x\}_{x \in \mathbb{Z}} \sim \mathbb{P}$, one has

- (3) Nearest neighbor Random Walk in Random Environment (RWRE)



$$\begin{aligned} \text{For fixing } \omega, \quad P_\omega(X_{n+1} = x+1 | X_n = x) &= \omega_x(1); \\ P_\omega(X_{n+1} = x-1 | X_n = x) &= \omega_x(-1). \end{aligned} \quad (2)$$

$P_\omega^\circ(\cdot)$ is called **quenched probability**;

$P^\circ(\cdot) := \int P_\omega^\circ(\cdot) \mathbb{P}(d\omega)$ is called **annealed probability**.



The history of the nearest neighbor RWRE:

- (1) 0-1 law and LLN (ω_i i.i.d.) (Solomon 1975 Ann. Prob.)
- (2) 0-1 law, LLN, CLT (ergodic ω)
(Alili 1999 JAP, Zeitouni 2004 LNM)

$$\text{Let } \rho_i = \frac{\omega_i(-1)}{\omega_i(1)}.$$

$\mathbb{E}(\log \rho_0)$ determines the recurrence and transience.

$\mathbb{E}(\rho_0)$ determines the LLN.

- (3) LDP (i.i.d. ω) (See Greven-Hollander 1994 Ann. Prob.).
- (4) LDP (ergodic ω)
(See Comets-Gantert-Zeitouni 2000 PTRF).
- (5) Kesten-Kozlov-Spitzer [6](1975) proved a stable limit theorem for the nearest neighbor RWRE.



Difficulties arise when randomizing ω

(1) Under the quenched probability P_ω^o , X_n cannot be written as sums of i.i.d. random variables any longer.
Simple random walk: $S_n = \sum_{i=1}^n \xi_i$, ξ_i i.i.d..

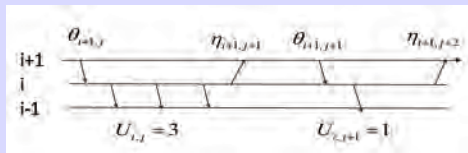
(2) Under the annealed probability P^o
 $\{X_n\}$ is not a Markov chain.

One needs new tools for further studies of RWRE.

“Branching Structure”: a powerful tool.



Branching structure for nearest neighbor RWRE



Define $T_n = \inf\{k > 0 : X_k = n\}$.

While $X_n \rightarrow \infty$, define

$$U_i^n = \#\{k \leq T_n : X_{k-1} = i, X_k = i-1\}.$$

$$P_\omega(U_i^n = k | U_{i+1}^n = 1) = [\omega_i(-1)]^k \omega_i(1).$$

Then

$$U_n^n = 0, U_{n-1}^n, \dots, U_1^n$$

are the first n generations of a branching process with immigration in random environment. Moreover,

$$T_n = n + 2 \sum_{i \leq n} U_i^n.$$



The function of the branching structure

The function of the branching structure:

- 1 One can count exactly the steps of the walk before T_n ,
$$T_n = n + 2 \sum_{i \leq n} U_i^n.$$
- 2 Give the explicit expression of “invariant density”, the key element of the method “The environment viewed from particle” (Zeitouni (2004) LNM).
- 3 Kesten-Kozlov-Spitzer (1975) proved the stable law for nearest neighbor RWRE.
- 4 Kesten (1977 Proceedings) proved the renewal theorem for nearest neighbor RWRE.
- 5 Gantert-Shi (2002 SPA) proved a limit theorem for the maximum of the occupation time.



Where this paper begins

Kesten-Kozlov-Spitzer (1975) (Branching structure for nearest neighbor RWRE)

In the literature we are aware of there is no description of the branching structure for non-nearest neighbor RWRE.

Questions:

(1) Are there any intrinsic connections between non-nearest neighbor RWRE and Branching Process in Random Environment?

(2) How to construct the branching structure for non-nearest neighbor RWRE? To what extent can we use the branching structure to study the limiting behaviors of non-nearest neighbor RWRE?



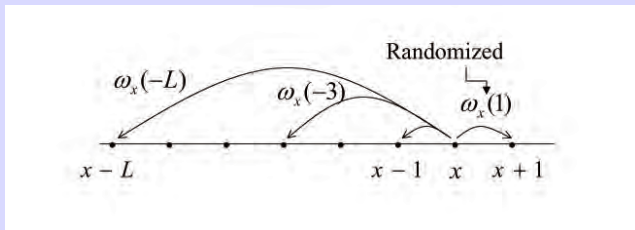
(L-1) RWRE

The model of the paper

(L-1) RWRE X_n :

$$P_\omega^o(X_{n+1} = x + l | X_n = x) = \omega_x(l),$$

$$\{\omega_x\}_{x \in \mathbb{Z}} = \{(\omega_x(-L), \dots, \omega_x(-1), \omega_x(1))\}_{x \in \mathbb{Z}} \sim \mathbb{P}.$$



The model

$\Lambda = \{-L, \dots, 1\} \setminus \{0\}$;

Σ : The simplex in \mathbb{R}^{L+1} .

$\Omega := \Sigma^{\mathbb{Z}}$, space of the environment.

μ : a probability measure on Σ .

$\mathbb{P} = \bigotimes^{\mathbb{Z}} \mu$ is the product measure on Ω , which makes $\omega_x, x \in \mathbb{Z}$ i.i.d..

Assume that \mathbb{P} satisfies some elliptic condition: There exists $\varepsilon > 0$, such that

$$\mathbb{P}(\omega_0(-l)/\omega_0(1) \geq \varepsilon, \forall l \in \Lambda) = 1. \quad (3)$$

For $\omega \in \Omega$, define (L-1) RWRE as a Markov chain, with initial value $X_0 = 0$, transition probabilities

$$P_\omega(X_{n+1} = x + z | X_n = x) = \omega_x(z), \forall x \in \mathbb{Z}, z \in \Lambda. \quad (4)$$

$P_\omega^o(\cdot)$: is called the quenched probability;

$P^o(\cdot) := \int_{\Omega} P_\omega^o(\cdot) \mathbb{P}(d\omega)$: is called the annealed probability.



(L-1) RWRE

(L-R) RWRE was first introduced in **Key (1984) Ann. Prob.**.

Brémont (2002) Ann. Prob. gave the 0-1 law and LLN.

Let

$$\overline{M}_i = \begin{pmatrix} a_i(1) \cdots a_i(L-1) a_i(L) \\ 1 & \cdots & 0 & 0 \\ \vdots & \ddots & \vdots & \vdots \\ 0 & \cdots & 1 & 0 \end{pmatrix}, \quad (5)$$

where $a_i(l) = \frac{\omega_i(-l) + \dots + \omega_i(-L)}{\omega_i(1)}$.

Let γ_L be the great Lyapunov exponent of $\{\overline{M}_i\}$.



Relations between certain matrices

Define

$$M_{-i} = \begin{pmatrix} b_{-i}(1) & \cdots & b_{-i}(L-1) & b_{-i}(L) \\ 1 + b_{-i}(1) & \cdots & b_{-i}(L-1) & b_{-i}(L) \\ \vdots & \ddots & \vdots & \vdots \\ b_{-i}(1) & \cdots & 1 + b_{-i}(L-1) & b_{-i}(L) \end{pmatrix}, \quad (6)$$

where $b_i(l) = \frac{\omega_i(-l)}{\omega_i(1)}$, $1 \leq l \leq L$;

$$B = \begin{pmatrix} 1 & & & \\ 1 & 1 & & \\ \vdots & \vdots & \ddots & \\ 1 & 1 & \cdots & 1 \end{pmatrix}. \quad (7)$$

Then $\overline{M}_i = B^{-1}M_iB$. Moreover, $\{\overline{M}_i\}$ and $\{M_i\}$ share the same great Lyapunov exponent γ_L .



Theorem (Brémont 2002 Ann. Prob.)

- (i) $\gamma_L < 0 \Rightarrow X_n \rightarrow \infty P^o\text{-a.s.};$
- (ii) $\gamma_L > 0 \Rightarrow X_n \rightarrow -\infty P^o\text{-a.s.};$
- (iii) $\gamma_L = 0 \Rightarrow -\infty = \liminf X_n < \limsup X_n = \infty P^o\text{-a.s..}$

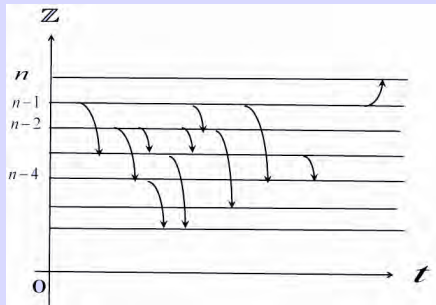
Also, Brémont showed a LLN under a condition called (IM) condition by him, that is,

$$\lim_{n \rightarrow \infty} \frac{X_n}{n} \rightarrow v P^o\text{-a.s..}$$



Branching structure for (L-1) RWRE

Branching structure for (L-1) RWRE ($\limsup_{n \rightarrow \infty} X_n = \infty$)



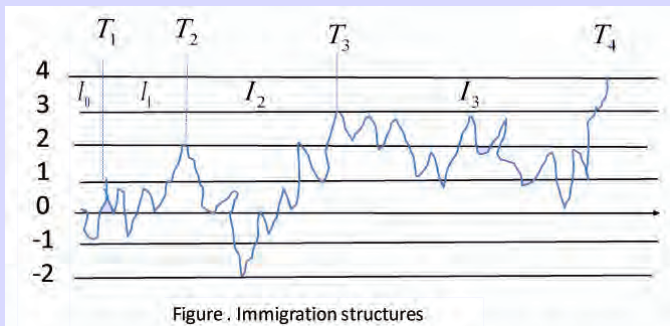
$$U_{i,l}^n = \#\{0 < k < T_n : X_{k-1} > i, X_k = i - l + 1\},$$

$$U_i^n := (U_{i,1}^n, U_{i,2}^n, \dots, U_{i,L}^n),$$

$$T_n = n + \sum_{i=-\infty}^{n-1} |U_i^n| + \sum_{i=-\infty}^{n-1} U_{i,1}^n = n + \sum_{i=-\infty}^{n-1} U_i^n(2, 1, \dots, 1)^T.$$



Path decomposition and immigration structure



This figure illustrates each of I_k corresponds to an immigrant and all immigration structures are independent by Markov property

Path decomposition

$$I_k =: \{X_m : T_k \leq m < T_{k+1}\}, \quad k = 0, 1, \dots, n-1.$$

For $1 \leq l \leq L$, $i < k$, define

$$U_l^n(k, i) = \#\{T_k \leq m < T_{k+1} : X_{m-1} > i, X_m = i - l + 1\}.$$

Then $U_l^n(k, i)$ records all steps by the walk from above i downwards to $i - l + 1$ in the random walk path I_k . Let

$$U^n(k, i) = (U_1^n(k, i), U_2^n(k, i), \dots, U_L^n(k, i)),$$

recording all steps by the walk crossing or reaching i from above i downwards in the random walk path I_k . For $i > k$, let $U^n(k, i) = \mathbf{0}$, and let $U^n(k, k) = e_1$. Then by definition of U_i^n and $U^n(k, i)$,

$$U_i^n = \sum_{k=(i+1) \vee 0}^{n-1} U^n(k, i).$$



Path decomposition

The offspring distribution of the type-1 particles

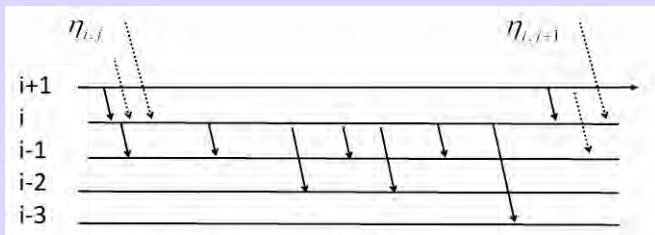


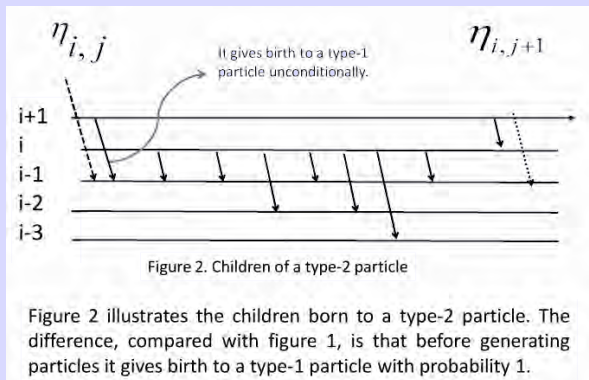
Figure 1. Children of a type-1 particle

This figure illustrates the children born to a type-1 particle. It has four type-1 children, two type-2 children, and one type-3 child.

$$\begin{aligned} P_{\omega}(U^n(k, i-1) = (u_1, \dots, u_L) | U^n(k, i) = e_1) \\ = \frac{(u_1 + \dots + u_L)!}{u_1! \dots u_L!} \omega_i(-1)^{u_1} \dots \omega_i(-L)^{u_L} \omega_i(1). \end{aligned}$$

Path decomposition

The offspring distribution of the type- $l (\geq 2)$ particles



For $2 \leq l \leq L$,

$$\begin{aligned} P_\omega \left(U^n(k, i-1) = (u_1, \dots, \mathbf{1} + u_{l-1}, \dots, u_L) \mid U^n(k, i) = e_l \right) \\ = \frac{(u_1 + \dots + u_L)!}{u_1! \dots u_L!} \omega_i(-1)^{u_1} \dots \omega_i(-L)^{u_L} \omega_i(1). \end{aligned}$$

Branching structure for (L-1) RWRE

Theorem 1 (Branching structure for (L-1) RWRE)

Suppose that $\limsup_{n \rightarrow \infty} X_n = \infty$. Let $x_0 = (2, 1, \dots, 1)$. Fix $n > 0$. Then

- (a) $T_n = n + \sum_{i=-\infty}^{n-1} |U_i^n| + \sum_{i=-\infty}^{n-1} U_{i,1}^n = n + \sum_{i=-\infty}^{n-1} U_i^n x_0$;
(b) For $1 \leq k \leq n-1$, $U^n(k, *)$ is a nonhomogeneous multitype branching process starting at time k , with branching mechanism

$$\begin{aligned} P_\omega(U^n(k, i-1) = (u_1, \dots, u_L) | U^n(k, i) = e_1) \\ = \frac{(u_1 + \dots + u_L)!}{u_1! \dots u_L!} \omega_i(-1)^{u_1} \dots \omega_i(-L)^{u_L} \omega_i(1), \end{aligned}$$

and for $2 \leq l \leq L$,

$$\begin{aligned} P_\omega(U^n(k, i-1) = (u_1, \dots, 1 + u_{l-1}, \dots, u_L) | U^n(k, i) = e_l) \\ = \frac{(u_1 + \dots + u_L)!}{u_1! \dots u_L!} \omega_i(-1)^{u_1} \dots \omega_i(-L)^{u_L} \omega_i(1). \end{aligned}$$



Theorem 1 (Continued)

further, conditioned on ω , $U^n(k, *)$, $k = 0, 1, \dots, n-1$ are i.i.d. and particles of $U^n(k, *)$ have independent lines of offspring.

$$U_{n-1}^n = \mathbf{0}, U_{n-2}^n, \dots, U_1^n, U_0^n, \quad (\text{Note: } U_i^n = \sum_{k=(i+1) \vee 0}^{n-1} U^n(k, i).)$$

are the first n generations of a multitype branching process with a type-1 immigration in each generation in random environment.

(c)

$$U_{n-1}^n = \mathbf{0}, U_{n-2}^n, \dots, U_1^n, U_0^n$$

shares the same distribution with the first n generations of a MBPIRE $\{Z_{-n}\}_{n \geq 0}$, where the branching mechanism of $\{Z_{-n}\}_{n \geq 0}$ will be given in (8) and (9).



Multitype branching process with immigration in random environment

For $k \in \mathbb{Z}$ define $Z(k, m)$ to be an L -type branching process starting at time k in random environment ω . That is, given ω ,

$$\begin{aligned}P_{\omega}(Z(k, m) = \mathbf{0}) &= 1, \text{ if } m > k, \\P_{\omega}(Z(k, k) = e_1) &= 1,\end{aligned}$$

and for $m < k$

$$\begin{aligned}P_{\omega}(Z(k, m) = (u_1, \dots, u_L) | Z(k, m+1) = e_1) \\= \frac{(u_1 + \dots + u_L)!}{u_1! \dots u_L!} \omega_{m+1}(-1)^{u_1} \dots \omega_{m+1}(-L)^{u_L} \omega_{m+1}(1),\end{aligned}\tag{8}$$

$$\begin{aligned}P_{\omega}(Z(k, m) = (u_1, \dots, u_{l-2}, u_{l-1} + 1, u_l, \dots, u_L) | Z(k, m+1) = e_l) \\= \frac{(u_1 + u_2 + \dots + u_L)!}{u_1! \dots u_L!} \omega_{m+1}(-1)^{u_1} \dots \omega_{m+1}(-L)^{u_L} \omega_{m+1}(1),\end{aligned}\tag{9}$$

$$l = 2, 3, \dots, L.$$



Multitype branching process with immigration in random environment

Let

$$Z_{-n} = \sum_{k=0}^{n-1} Z(-k, -n), \quad n > 0. \quad (10)$$

$\{Z_{-n}\}_{n \geq 0}$ is called **Multitype Branching Process with Immigration in Random Environment (MBPIRE)**.

Given ω , let M_{-i} be such a matrix whose l -th row is the expected value of particles born to a type- l particle at time $-i$. Then

$$M_{-i} = \begin{pmatrix} b_{-i}(1) & \cdots & b_{-i}(L-1) & b_{-i}(L) \\ 1 + b_{-i}(1) & \cdots & b_{-i}(L-1) & b_{-i}(L) \\ \vdots & \ddots & \vdots & \vdots \\ b_{-i}(1) & \cdots & 1 + b_{-i}(L-1) & b_{-i}(L) \end{pmatrix}, \quad (11)$$

where $b_i(l) = \frac{\omega_i(-l)}{\omega_i(1)}$, $1 \leq l \leq L$.



Remark 1 (Ideas for the branching structure)

for (L-1) RWRE, the case becomes much more complicated because there are overlaps between different jumps. So to construct the branching structure, we use the following new ideas:

- 1) We treat the jumps reaching or crossing i downward from above i as the particles of i ;
- 2) The types are determined by the position that the jumps reach. That is, if a jump crosses i downward from above i and reaches $i - l + 1$ at last, the jump is a type- l particle of i ;
- 3) We decompose the random walk path before T_n into n independent and non-intersection pieces and every piece corresponds to an immigration structure.



The first simple application

One follows from the branching structure immediately the following interesting result.

That is, for transient walk ($X_n \rightarrow \infty$), the tail of the minimum of the walk before hitting the positive real line decays faster than exponential rate.

Corollary 1

Suppose that $\gamma_L < 0$, and there is $q > 0$ such that $\mathbb{E}\|M_0\|^q < \infty$. Then for $t > 0$, one has that

$$P\left(\min_{0 \leq n \leq T_1} X_n < -t\right) < k_1 e^{-k_2 t}$$

where k_1 and k_2 are positive constants.



Testing of the branching structure

Consider a special (2-1) random walk. Fix $\omega := (\dots, \omega_0, \omega_0, \omega_0, \dots) \in \Omega$, $\omega_0 = (q_1, q_2, p)$, where $p, q_1, q_2 \geq 0$, $p + q_1 + q_2 = 1$.

Suppose $\{X_n\}_{n \geq 0}$ is a random walk with $X_0 = 0$ and transition probabilities

$$P_\omega(X_{n+1} = j | X_n = i) = \begin{cases} p & \text{if } j = i + 1, \\ q_1 & \text{if } j = i - 1, \\ q_2 & \text{if } j = i - 2. \end{cases}$$

Theorem 2 (The mean of T_1)

Suppose that $p > q_1 + 2q_2$. Then P_ω -a.s., $T_1 < \infty$ and $E_\omega(T_1) = \frac{1}{p - q_1 - 2q_2}$.

Proof. (Method 1: By Ward equation)

$$1 = E_\omega(X_{T_1}) = E_\omega(T_1)E_\omega(X_1) = (p - q_1 - 2q_2)E_\omega(T_1).$$

(Method 2: Using branching structure) $T_1 = 1 + \sum_{i=-\infty}^0 U_i^1(2, 1)^T$.



Stable Law for (L-1) RWRE

Let $\rho = \frac{1-\omega_0(1)}{\omega_0(1)}$.

Condition (C)

(C1) $\mathbb{E}(\log^+ \rho) < \infty$, where $\log^+ x := 0 \vee \log x$.

(C2) $\mathbb{P}(\rho > 1) > 0$.

Under (C2), there exists $\kappa_0 > 2$ such that

$$\mathbb{E} \left[\left(\min_{1 \leq i \leq L} \left\{ \sum_{j=1}^L M_0(i, j) \right\} \right)^{\kappa_0} \right] = \mathbb{E}(\rho^{\kappa_0}) > 1. \quad (12)$$

Fixing κ_0 in (12), we give a new condition

(C3) $\mathbb{E}(\rho^{\kappa_0} \log^+ \rho) < \infty$.

Let ϱ be the greatest eigenvalue of M_0 . Then $\varrho > 0$.

(C4) The group generated by $\text{supp}[\log \varrho]$, is dense in \mathbb{R} . \square

Kesten 1973 proved that if $\gamma_L < 0$, then there is unique $\kappa \in (0, \kappa_0]$ such that

$$\log \rho(\kappa) := \lim_{n \rightarrow \infty} \frac{1}{n} \log \mathbb{E}(\| M_0 M_{-1} \cdots M_{-n+1} \|^\kappa) = 0. \quad (13)$$



A Stable Law for (L-1) RWRE

Theorem 3 (Stable law for (L-1) RWRE)

Suppose that Condition **(C)** holds and $\gamma_L < 0$. Fix κ in (13). Let $L_\kappa(x)$ be a κ -stable law (if $\kappa < 1$, L_κ has support $[0, \infty)$; if $\kappa > 1$, L_κ has zero mean).

(i) if $0 < \kappa < 1$, then

$$\lim_{n \rightarrow \infty} P(n^{-\frac{1}{\kappa}} T_n \leq x) = L_\kappa(x),$$

$$\lim_{n \rightarrow \infty} P(n^{-\kappa} X_n \leq x) = 1 - L_\kappa(x^{-\frac{1}{\kappa}});$$

(ii) if $\kappa = 1$, then for suitable $D(n) \sim \log n$ and $\delta(n) \sim (A_1 \log n)^{-1} n$,

$$\lim_{n \rightarrow \infty} P(n^{-1}(T_n - A_1 n D(n\mu^{-1})) \leq x) = L_1(x),$$

$$\lim_{n \rightarrow \infty} P(n^{-1}(\log n)^2(X_n - \delta(n)) \leq x) = 1 - L_1(-A_1^2 x);$$



Theorem 3 (Stable law for (L-1) RWRE (Continued))

(iii) if $1 < \kappa < 2$, then

$$\lim_{n \rightarrow \infty} P \left(n^{-\frac{1}{\kappa}} (T_n - A_\kappa n) \leq x \right) = L_\kappa(x),$$

$$\lim_{n \rightarrow \infty} P \left(n^{-\frac{1}{\kappa}} \left(X_n - \frac{n}{A_\kappa} \right) \leq x \right) = 1 - L_\kappa(-x A_\kappa^{1+\kappa^{-1}});$$

(iv) if $\kappa = 2$, then

$$\lim_{n \rightarrow \infty} P \left(\frac{T_n - A_2 n}{B_1 \sqrt{n \log n}} \leq x \right) = \Phi(x),$$

$$\lim_{n \rightarrow \infty} P \left(A_2^{\frac{3}{2}} B_1^{-1} (n \log n)^{-\frac{1}{2}} \left(X_n - \frac{n}{A_2} \right) \leq x \right) = \Phi(x);$$

Theorem 3 (Stable law for (L-1) RWRE (Continued))

(v) if $\kappa > 2$, then

$$\lim_{n \rightarrow \infty} P\left(\frac{T_n - B_3 n}{B_2 \sqrt{n}} \leq x\right) = \Phi(x),$$

$$\lim_{n \rightarrow \infty} P\left(B_3^{\frac{3}{2}} B_2^{-1} n^{-\frac{1}{2}} \left(X_n - \frac{n}{B_3}\right) \leq x\right) = \Phi(x),$$

where $0 < A_\kappa, B_i < \infty$ are suitable constants, and $\Phi(x) := \frac{1}{\sqrt{2\pi}} \int_{-\infty}^x e^{-\frac{s^2}{2}} ds$.

Ideas of the proof: $T_n = n + \sum_{i=-\infty}^{n-1} U_i^n x_0$. It suffices to show that $\sum_{i=0}^{n-1} U_i^n x_0$ converges to L_κ after suitable renormalization. But

$$\sum_{i=0}^{n-1} U_i^n x_0 \stackrel{\mathcal{D}}{=} \sum_{i=0}^{n-1} Z_{-i} x_0 \approx \sum_{k=1}^{n/E(\nu)} W_k x_0.$$



The slowdown properties of RWRE

From (i) of Theorem 3 (Stable Law) one has the following theorem, which tells that the walk slows down after randomizing the environment.

Theorem 4 (The slowdown properties of RWRE)

Suppose that Condition **(C)** holds, $\gamma_L < 0$, and that $0 < \kappa < 1$. Then for $\kappa < s < 1$ one has that

$$\lim_{n \rightarrow \infty} \frac{X_n}{n^s} = 0, \quad P^o\text{-a.s..}$$

For simple random walk (linear speed)

$$S_n \rightarrow \infty \Leftrightarrow \frac{S_n}{n} \rightarrow E(\xi_1) > 0.$$

But Theorem 4 reveals that for (L-1) RWRE (sub-linear speed),

$$\text{though } X_n \rightarrow \infty, \quad \frac{X_n}{n^s} \rightarrow 0, \quad \kappa < s.$$



LLN for (L-1) RWRE

For $n \geq 0$, define $\bar{\omega}(n) = \theta^{X_n} \omega$. Then $\{\bar{\omega}(n)\}$ is a Markov chain with transition kernel

$$\bar{P}(\omega, d\omega') = \omega_0(1) \delta_{\theta\omega=\omega'} + \sum_{l=1}^L \omega_0(-l) \delta_{\theta^{-l}\omega=\omega'}.$$

Define

$$\pi(\omega) := \frac{1}{\omega_0(1)} \left(1 + \sum_{i=1}^{\infty} e_1 \bar{M}_i \cdots \bar{M}_1 e_1^T \right).$$

Let $\tilde{\pi}(\omega) = \frac{\pi(\omega)}{\mathbb{E}(\pi(\omega))}$. Then one has



Theorem 5 (LLN and Invariant Measure)

Suppose that $\mathbb{E}(\pi(\omega)) < \infty$. Then

(i) $\gamma_L < 0$;

(ii) $\tilde{\pi}(\omega)\mathbb{P}(d\omega)$ is invariant under kernel $\bar{P}(\omega, d\omega')$, that is

$$\int 1_B \tilde{\pi}(\omega)\mathbb{P}(d\omega) = \iint 1_{\omega' \in B} \bar{P}(\omega, d\omega') \tilde{\pi}(\omega)\mathbb{P}(d\omega), \quad B \in \mathcal{F};$$

(iii) moreover \mathbb{P} -a.s., $\lim_{n \rightarrow \infty} \frac{X_n}{n} = \frac{1}{\mathbb{E}(\pi(\omega))}$.

Remark 2

LLN was proved in Brémont (2002) where the author introduced a so-called **(IM) Condition**, and gave the Invariant Measure by studying its transition probability. By the branching structure we can calculate the Quenched Mean $E_\omega^o(T_1)$ to give the explicit expression of the Invariant Measure, and prove the LLN directly by a method known as “the environment viewed from particles”.

The tail of the expected value of the total progeny of a immigration

Let $\{Z_{-n}\}_{n \geq 0}$ be the MBPIRE with negative time defined in (10). Let

$$Y_{-k} = \sum_{m=k+1}^{\infty} Z(-k, -m).$$

Note that for $m > k$, one has $E_{\omega}(Z(-k, -m)) = M_{-k}M_{-k-1} \cdots M_{-m+1}$. Then

$$\eta_{-k} := \sum_{m=k+1}^{\infty} M_{-k} \cdots M_{-m+1} \quad (14)$$

the expected offspring matrix of Y_{-k} . We prove the follow theorem:



Total progeny of an immigration

Theorem 6 (Tail of total progeny of an immigration)

Suppose that Condition **(C)** holds and $\gamma_L < 0$. Then for κ in (13) and suitable constant $K_2 = K_2(x_0) \in (0, \infty)$, one has that

$$\lim_{t \rightarrow \infty} t^\kappa \mathbb{P}(x\eta_0x_0 \geq t) = K_2|xB|^\kappa, \quad (15)$$

where $x \in \mathbb{R}^L$ with all coordinates positive and $|x| > 0$.

Remark 3

Indeed, Kesten (1973) showed that

$$\lim_{t \rightarrow \infty} t^\kappa \mathbb{P}(x\eta_0x_0 \geq t) = K(x, x_0),$$

where K depends on x and x_0 . The main contribution of Theorem 6 is that it tells how the constant K depends on x . K_2 in (15) is independent of x .



Tail of the total population before regeneration

Let $\nu_0 \equiv 0$, and define recursively

$$\nu_n = \min\{m > \nu_{n-1} : Z_{-m} = \mathbf{0}\}, \quad \forall n > 0.$$

Write ν_1 simply as ν .

ν is the regeneration time of $\{Z_{-n}\}_{n \geq 0}$;

From the point of view of the random walk ν is $\{X_n\}_{n \geq 0}$ the first position where the walk will never revisit after passing it.

Define

$$W = \sum_{k=0}^{\nu-1} Z_{-k}.$$

W is the total population born before regeneration time ν .



Tail of the total population before regeneration

Theorem 7 (Tail of W)

Suppose that Condition **(C)** holds and $\gamma_L < 0$. If $\kappa > 2$, $E((Wx_0)^2) < \infty$; if $\kappa \leq 2$, there exists $0 < K_3 < \infty$ such that

$$\lim_{t \rightarrow \infty} t^\kappa P(Wx_0 \geq t) = K_3.$$

Remark 4

Theorem 7 reveals that the total population of the MBPIRE before regeneration, Wx_0 , belongs to the domain of attraction of a κ stable law. Moreover Theorem 7 is the key step for the proof of the stable limit theorem (Theorem 3.).



Reference

- [1] Alili, S. (1999). Asymptotic behavior for random walks in random environments. *J. Appl. prob.* 36, 334-349.
- [2] Brémont, J. (2002). On some random walks on \mathbb{Z} in random medium. *Ann. of Probab.* 3, 1266-1312.
- [3] Comets, F., Gantert, N., Zeitouni, O. (2000). Quenched, annealed and functional large deviations for one-dimensional random walk in random environment. *Prob. Theory Rel. Fields* 118, 65-114.
- [4] Greven, Andreas, den Hollander, Frank. (1994). Large deviations for a random walk in random environment. *Ann. Prob.* 22 , 3, 1381-1428.
- [5] Hong,W.M. and Wang, H.M. (2010). Quenched moderate deviations principle for random walk in random environment. *Science in China, Series A, To appear.*



Reference

- [6] Kesten, H., Kozlov, M.V., Spitzer, F. (1975). A limit law for random walk in a random environment. *Compositio Math.* 30, pp. 145-168.
- [7] Kesten, H. (1977). A renewal theorem for random walk in a random environment. *Proceedings of Symposia in Pure Mathematics*, Vol. 31, pp. 67-77.
- [8] Key, E.S. (1984). Recurrence and transience criteria for random walk in a random environment. *Ann. Prob.* 12, pp. 529-560.
- [9] Gantert, N. and Shi, Z. (2002). Many visits to a single site by a transient random walk in random environment. *Stoc. Proc. Appl.* 99, pp. 159-176.
- [10] Solomon, F. (1975). Random walks in random environments. *Ann. prob.* 3, 1-31.
- [11] Zeitouni, O. (2004). Random walks in random environment. *LNM 1837, J. Picard (Ed.), 189-312, Springer-Verlag Berlin Heidelberg. (2004).*



非常感谢
Thanks a lot!

